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## Proximity and Horizontal Policies: The Backbone of Export Diversification and Complexity

by Gonzalo Salinas

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I N T E R N A T I O N A L M O N E T A R Y F U N D

**IMF Working Paper**

Western Hemisphere Department

**Proximity and Horizontal Policies:**

**The Backbone of Export Diversification and Complexity**

**Prepared by Gonzalo Salinas<sup>1</sup>**

Authorized for distribution by Sònia Muñoz

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**Abstract**

The lack of a clear link between general economic fundamentals and export diversification indicators in the literature has fueled the belief that industrial policies are an absolute requisite to diversify exports. This paper, however, does find a strong statistical connection between horizontal policies and diversification by making two novel changes to traditional methodologies: using export categories that lead to diversification (for example, manufactures) as dependent variables, and using a gravity-equation regression setting. Proximity to other economies explains about a third of cross-country heterogeneity in targeted exports, and four fifths together with horizontal policies. Australia, Chile, and New Zealand emerge as new role models for diversification policies.

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## I. INTRODUCTION

Export diversification is a commonly sought objective in emerging and developing economies for several reasons. Numerous studies indicate that a more diversified export basket is statistically associated with lower output volatility (Haddad and others, 2013), and that the latter is associated with higher long-term output growth (Ramey and Ramey, 1995; Hnatkovska and Loayza, 2004). This is of outmost importance to most developing countries, as their export baskets tend to be concentrated on a handful of raw commodities with highly volatile international prices. Moreover, production of raw commodities is not labor intensive, and therefore it is argued that nurturing other more labor-intensive export sectors could improve labor market outcomes and socioeconomic conditions.

Some economists further suggest that there are some *superior* products that countries should export because they are expected to have higher economic payoff in different ways. Hallak (2006) classifies export products according to their *quality* and shows that advanced economies tend to import more of those higher quality products. Products are classified according to their *sophistication* in Hausmann and others (2006), and *complexity* in Hidalgo and Hausmann (2009), and these studies suggest that having more sophisticated/complex exports is associated with higher future economic growth.

While there are several studies aiming to identify policies that lead to export diversification and the development of *superior* exports, there are important caveats in their methodological approaches. First, the dependent variables commonly used are indices (such as concentration, sophistication, or complexity indices) that are greatly affected by exogenous changes in international commodity markets, which weakens their link to policy variables. Second, the explanatory variables included are commonly limited to those associated with higher productivity (for example, education, institutions), neglecting the importance of labor costs and gravity equation-related variables for exports development (especially of manufacturing products) according to standard trade models.

This paper therefore makes two fundamental shifts in the analysis of export diversification and its policy determinants. Instead of using diversification/concentration, sophistication, or complexity indices as dependent variables, it uses levels of export categories (non-hydrocarbon/mineral (NHM), manufacturing, services, complex) that can diversify commodity dependent export baskets.<sup>2</sup> In addition, following a standard Eaton and Kortum (2002) Ricardian model, it adds gravity equation and labor costs variables to the commonly included productivity-related independent variables. Income per capita is introduced as a control for endogeneity and removed for goodness of fit estimation.

The proposed methodological changes allow for a substantial estimation improvement as coefficients are more robustly significant and goodness of fit is higher than in other studies.

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<sup>2</sup> In this paper, the term *NHM exports* includes only goods not services. Service exports are identified separately for analytical purposes as their dependence on distance and economic policies is considerably different from that of NHM goods exports.

The analysis finds a strong relation between gravity related variables and NHM, manufacturing, complex, and service exports. It also finds that governance, educational attainment, infrastructure quality, and trade policy openness are robustly related to these exports. Adding these horizontal policies to countries' proximity to international markets explains above 80 percent of cross-country variation in the targeted exports.

The paper first discusses in more detail, in Section II, the justification for the two analytical shifts proposed above, while Section III delineates this methodological approach. Section IV describes the regional and country level variation in NHM exports, as well as in independent variables. Econometric analysis in Section V statistically identifies the factors that are associated with the development of targeted exports and their explanatory power. Section VI discusses the main policy implications of these statistical findings.

## II. THEORETICAL AND EMPIRICAL CONSIDERATIONS

### *The Dependent Variable*

Most empirical attempts to identify the factors that foster export diversification use as dependent variable an export concentration index, such as the Herfindahl-Hirschman Index (HHI), while those aiming to identify the determinants of *superior* exports use several indices of exports *superiority*, such as the Exports Sophistication Index (ESI) (Hausmann and others, 2005; Weldemicael, 2012) or the Economic Complexity Index (ECI) (Hidalgo and Hausmann, 2009).

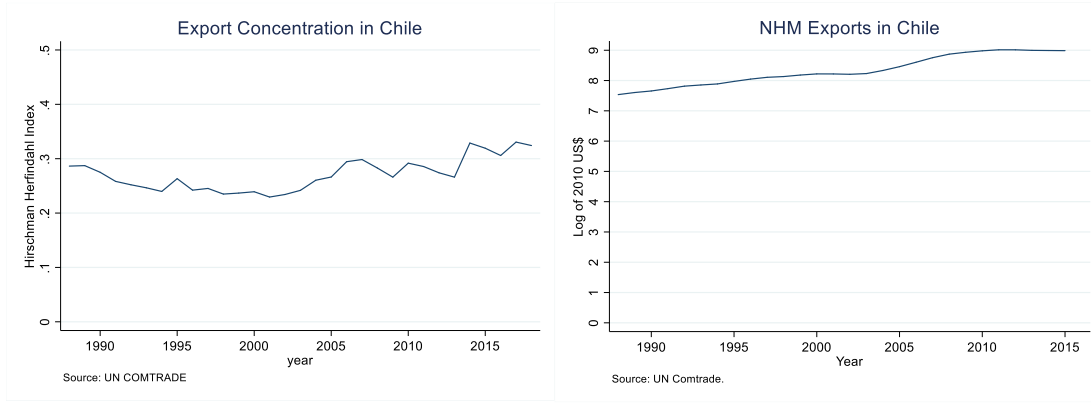
Nevertheless, these indices are substantially affected by exogenous factors, thus weakening their statistical link to policy determinants. Take for instance the HHI of export concentration for country  $j$  including exports ( $x$ ) of several sectors ( $s$ ):

$$(1) HHI_j = \sum_s \left( \frac{x_{sj}}{\sum_s x_{sj}} \right)^2$$

This index is higher when the nominal export value of one or few commodities is high relative to the total export basket, indicating more (less) exports concentration (diversification). In most developing countries, partly due to their weak production capacity, a handful of hydrocarbon/mineral (HM) exports account for most of their total exports. Hence when aiming to diversify exports these countries seek policies to nurture NHM products. If successful, the value of these products will narrow the gap with respect to the dominant HM exports and this would reduce their HHI.

But the HHI can also significantly fluctuate in response to variations in the nominal value of their HM exports, which are commonly the result of largely exogenous events such as changes in international commodity prices or findings of additional HM reserves. Such fluctuations can considerably weaken the statistical relationship between policy frameworks and the targeted development of NHM exports needed to diversify export baskets.

This is quite evident when looking at the evolution of the concentration index in a commodity exporting country. For instance, Chile's HHI in the early 2000s markedly reverted its previously downward trend in the absence of any substantial reform to its policy framework (see figures below). This surge in concentration is most evidently related to the international copper boom, which multiplied the value of Chile's copper exports from US\$ 8 billion in 2003 to a peak of US\$ 54 billion in 2011, when it accounted for half of its goods exports. The continued growth in per capita NHM exports during that period confirms that the surge in export concentration was not related to a weakening in Chile's NHM export potential. Because most countries that seek export diversification are strongly dependent on HM exports, this disconnect between the HHI and policy determinants of NHM exports due to commodity fluctuations is highly consequential.



A similar complication occurs when trying to identify a statistical relation between policy variables and *superior* exports by using sophistication or complexity indices as dependent variables. These indices are broadly a product of the sophistication/complexity of each exported product times the product's share in the country's export basket. For instance, the ESI of country  $j$  is calculated as the sum of the sophistication ( $PRODY$ ) of each exported product weighted by its share in its total exports:

$$(2) ESI_j = \sum_s \left( \frac{x_{js}}{\sum x_{js}} \right) PRODY_s$$

In turn, product sophistication is calculated as the weighted average of GDP per capita of the countries exporting product  $s$ , where the weight is the value-share of product  $s$  in country's  $j$  overall export basket divided by the value-shares across all countries exporting that good:

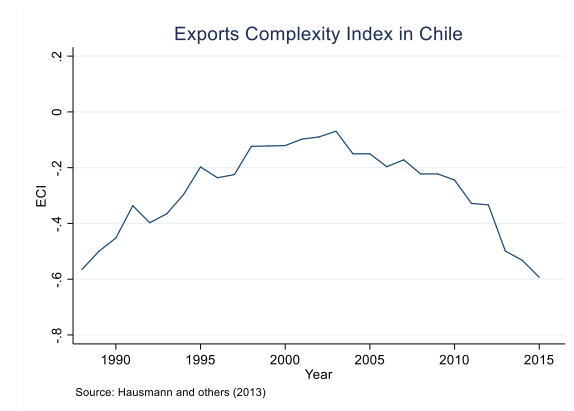
$$(3) PRODY_s = \sum_j \frac{\left( \frac{x_{js}}{\sum_i x_{js}} \right)}{\left( \sum_j \left( \frac{x_{js}}{\sum_i x_{js}} \right) \right)} GDPpc_j$$

Hence, a product is estimated as more sophisticated when it is mainly produced in higher income countries. Because HM exports dominate the export baskets of lower GDP per capita countries, in general, HM exports are estimated as less sophisticated than NHM exports:

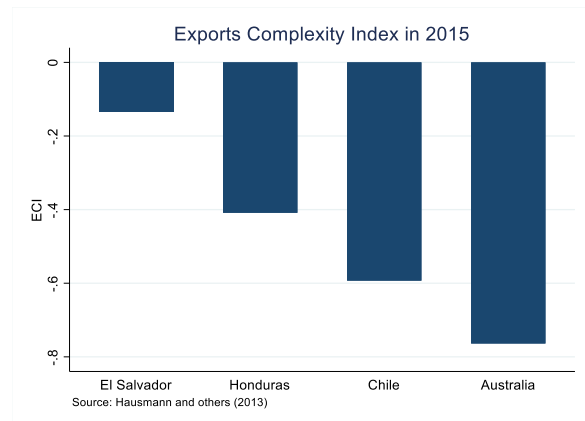
$$(4) PRODY_{HM} < PRODY_{NHM}$$

Since exports ( $x$ ) are measured in nominal values, exogenous increases in international HM prices or HM discoveries lower the ESI without any change in sophisticated exports. Regression specifications that aim to establish a link between policies and sophisticated exports are thus weakened by exogenous commodity related fluctuations. A similar analysis applies for the ECI<sup>3</sup>.

Chile during the early 2000s is also an illustrative case of how these indices can mislead the identification of policies that foster *superior* exports. Chile's ECI plummeted from close to zero in 2000 to -0.6 in 2015, a considerable fall as the ECI broadly ranges between -2.5 and 2.5. This decline seems at odds with the sustained productivity growth that Chile experienced those years which, a priori, should have increased its capacity to produce complex goods for exporting. As was the case with the HHI, Chile's ECI decline is most evidently related to the boom of copper (a low complexity product), thus showing how commodity fluctuations erode the relation between target variable (complex exports) and policy variables.<sup>4</sup>



The disconnect between the ECI and a country's policy framework is similarly evident in cross-country comparisons. A priori, the advanced Australian economy, with strong institutional and educational quality, should be more capable of producing complex products than Latin American countries. Yet, for example, the ECI of Australia is considerably below the ECIs of El Salvador and Honduras. According to its authors the ECI is a proxy for *productive capabilities* and



<sup>3</sup> The ECI of a country is calculated based on the diversity of exports a country produces and their ubiquity, or the number of countries able to produce them (and those countries' complexity). This index aims to measure the knowledge in a society in terms of the products it exports (Hausmann and others, 2013), but this is questionably the case as it is substantially affected by commodity fluctuations that are not related to the knowledge or productivity of economic agents.

<sup>4</sup> As an example of a similar disconnect in oil exporting countries, Nigeria's ECI has considerably deteriorated during oil price booms (in the early 1970s and early 2000s) and improved significantly in 2008, as a result of the oil price collapse of that year. At a regional level, as noted in Ding and Hadzi-Vaskov (2017), a growing trend in the share of complex exports in Latin American and Caribbean in the 1990s was reversed in the 2000s because of the commodity price boom, as the region is a major exporter of these products.

measures the *knowledge of a society* (Hausmann and others, 2013), but it is questionable that Australia's *productive capabilities* are inferior in this illustrative cross-country comparison. Australia's low ECI is likely related to its exogenously high mineral endowment and consequent high exports of minerals, which are low complexity products.

Another illustrative case of the limitations of the ECI as a measure of complexity due to natural resource abundance is the U.S. state of Texas. Despite being a global technology leader its ECI is only 0.29, similar to the Philippines. This evident inconsistency likely results from Texas superlative petroleum endowments and the extremely low (-2.57) Product Complexity Index (PCI) of Petroleum Oils in Hausmann and others (2013).

This dependency of the ECI on exogenous commodity developments is systemic across countries. Fixed effect regressions including most countries (Table A.1) indicate that the ECI is strongly associated with resource wealth as defined in Sachs and Warner (1995). It is thus likely that the ECI's statistical relation with future GDP growth described in Hidalgo and Hausmann (2009) is related to the *resource curse* identified earlier in several studies including in Sachs and Warner (1995).

The evident disconnect between the above discussed indices and policy determinants that foster diversification and export *superiority* can be effectively addressed by focusing directly on the evolution of the export products that lead to diversification or export *superiority*. Since export diversification is commonly sought in countries that are dependent on a handful of HM exports, the relevant dependent variable is the value of NHM, manufacturing, and/or services exports. And when aiming to foster *superior* exports the dependent variable can be directly defined as the value of those *superior* exports itself. These dependent variables can be normalized by population or labor force to control for size.

An alternative solution to deal with the weakened link between the HHI and diversification policy determinants would be to remove HM exports from the calculation of the HHI. In practice, however, there is hardly any country that has an export basket concentrated on few NHM products and therefore strongly needs to diversify its NHM export basket. A statistical analysis of HHI excluding NHM exports as dependent variable would thus have little practical relevance. A similar option, to calculate an ECI excluding HM exports to measure a country's complexity, would not accurately measure the overall *knowledge of society* in cases in which a small, non-representative elite is capable of exporting complex products but most of society does not export and is involved in low productivity activities.<sup>5</sup>

Fixed effects regressions in Table A.1 and the correlation table below show that complex exports per capita is positively associated with the ECI. Central to the analysis in this paper, this matrix also shows that the ratio of complex exports per capita is not significantly

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<sup>5</sup> One can think of many developing countries in which a very small elite group produces complex exports, but the export basket is dominated by HM products and most of the labor force is involved in a low-productivity informal economy. Such country would have a high ECI excluding HM exports, but this would be a poor measure of the overall *knowledge of society*. On the other hand, the country's level of complex exports per capita would surely capture the weak *productive capabilities* of its overall economy.



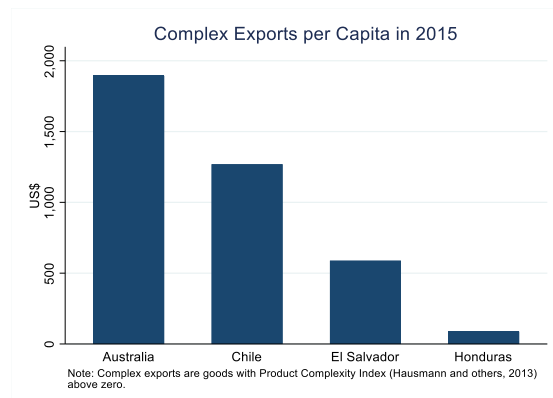
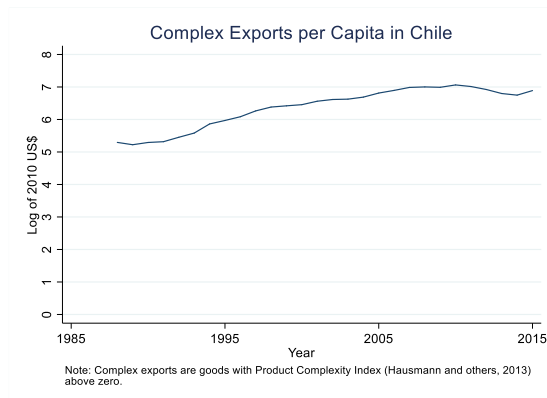
correlated to the ratio of HM exports to GDP. The ratio of complex exports per capita, as intended, is not affected by natural resource abundance and commodity price fluctuations, which should improve the identification of the factors that foster the development of complex exports.

**Correlation ECI, natural resource abundance, and complex exports per capita**

	ECI	ECI Plus	HM exports to GDP	Complex Exports Per Capita (Log)
ECI	1.00			
ECI Plus	0.91	1.00		
HM exports to GDP	-0.23	-0.17	1.00	
Log Complex Exports Per Capita	0.76	0.76	0.07	1.00

Source: Hausmann and others (2013) and author's calculations.

The analysis substantially changes when focusing directly on the evolution of the targeted export groups per capita. As seen above, although Chile's HHI pointed to declining diversification in the early 2000s, NHM exports per capita continued to increase during that period. The picture similarly changes when looking directly at the value of per capita *complex exports*, which can be defined, for instance, as those in the upper half of the PCI (with PCI above zero). Unlike the ECI, the real value of Chile's complex exports per capita continued to grow during the copper boom and, as expected, complex exports per capita is higher in Australia than in Honduras and El Salvador.



### *Independent Variables*

Another fundamental shift needed to sharpen the identification of factors that foster export diversification and *superior* exports is basing regression specifications on standard trade theory. This second shift is partly related to the first one, as the dependent variables proposed are levels of exports (not indices) and these have been extensively modelled in the international trade literature.

The analysis can thus be broadly rooted, for example, on a standard EK02 (Eaton and Kortum, 2002) Ricardian general equilibrium model, which takes into accounts distance and labor costs in addition to productivity. We can relate the target export categories (NHM, manufacturing, complex, and services) to the manufacture sector in EK02's two-sector setting of manufactures and non-manufactures:<sup>6</sup>

$$(5) \quad \frac{X_{ni}}{X_n} = T_i \left( \frac{\gamma d_{ni} w_i^\beta p_i^{1-\beta}}{p_n} \right)^{-\theta}$$

where the fraction of total expenditure of country  $n$  on manufacturing goods from country  $i$  ( $X_{ni}$ ) divided by its total expenditure ( $X_n$ ), is a function of country  $i$ 's state of technology ( $T_i$ ), wages in country  $i$  ( $w_i$ ), and prices in both countries.<sup>7</sup> Note that while distance-related variables are mostly exogenous, those related to technology and wages are largely determined by public policies of the exporting economy.

Other empirical studies on the determinants of export diversification and *superior* exports include independent variables that are related to productivity/technology (T-variables) of the exporting country ( $i$ ), but do not include wage and gravity-related variables. This omission implicitly assumes that labor costs and gravity-equation variables do not have a different effect in the targeted export groups (NHM or *superior* exports) from the non-targeted exports (homogeneous elasticities across exports types). A significant exception is Weldemicael (2012), which finds that distance to markets is strongly related to the ESI, thus implying a significant difference in the impact of distance on more sophisticated exports from less sophisticated exports. In line with this finding, the statistical analysis below identifies heterogenous elasticity-to-distance across several export groups.

Empirical findings from the GVC literature also hint at the importance of distance to large markets and other gravity equation variables in the development of *superior* exports. Raei and others (2019) and Kowalski and others (2015) identify gravity variables as key determinants of Global Value Chain (GVC) participation. Since participation in GVCs is seen as a major force behind the growth of more complex, manufacturing products, it is very likely that gravity-related variables are significant determinants of *export superiority*.

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<sup>6</sup> Note though that since the EKO2 and a large class of international trade models generate isomorphic gravity equations (Arkolakis and others, 2012), the results of our analysis should be broadly robust to model selection.

<sup>7</sup> Parameter  $\gamma$  is a measure of the sensitivity of local prices to foreign cost structures and geographic barriers.  $\theta$  represents product homogeneity across countries, which governs comparative advantage. A low  $\theta$  implies high product variability and in that case comparative advantage exerts a bigger force for trade.  $\beta$  is labor's share in production, while  $(1 - \beta)$  is intermediate inputs' share in production.

### III. METHODOLOGY AND DATA

Part of the analysis in this paper focuses on the value of each export category normalized by population to allow cross-country comparisons among exporters. Admittedly, NHM exports per capita are not totally immune to commodity booms and natural resource abundance. Commodity booms could trigger Dutch Disease dynamics attracting factors of production towards HM export sectors and away from NHM exports. As seen in the case of Chile in the early 2000s and as in the stylized facts charts presented below, this effect does not appear to be substantial within a country and across time. However, exports of HM products, for sure, use up significant labor, directly and indirectly, thereby reducing NHM exports per capita relative to non-commodity exporters. It is worth keeping this in mind especially in cross-country comparisons of NHM exports per capita.

Normalizing by population also makes the dependent variable less sensitive to commodity fluctuations than alternative normalizations. Normalizing NHM exports by dividing them by GDP is ruled out as GDP in HM abundant countries by definition includes HM exports. Dividing NHM exports by GDP excluding HM sectors is similarly rejected because commodity boosts and busts indirectly affect other sectors, such as commerce and services.

A potential drawback of normalizing targeted exports through dividing by population is that such ratio can be exogenously affected by the share of the population in working age, which can affect cross-country and cross-time comparisons. This could hence weaken the link between policies and targeted exports and suggest normalization by the labor force could be preferable. Nonetheless, participation in the labor force can be endogenous to economic activity and therefore to policy determinants under consideration in this study. For example, if policy determinants are effective in fostering the targeted exports partly by increasing labor force participation, the exports-to-labor force ratio would be relatively unchanged and not reflect the success of those policies. In any case, the identification of policy determinants of diversification does not require normalization of the dependent variable and while the analysis on goodness of fit does require normalization, it is largely robust to normalization by labor force instead of population.

A preliminary panel econometric specification can be derived from equation (1), rearranging its terms, adding the time dimension, and log linearizing it:<sup>8</sup>

$$(6) \quad X_{ni,t} = \alpha_1 \log(\gamma) + \alpha_2 \log(X_{n,t}) + \alpha_3 \log(d_{ni}) + \alpha_4 \log T_{i,t} + \alpha_5 \log(w_{i,t}) + \vartheta_{ni} + \mu_t + \varepsilon_{ni,t}$$

where year ( $\mu_t$ ) and country-pair fixed effects ( $\vartheta_{ni}$ ) are introduced. In addition to distance between countries itself, other commonly used gravity equation variables (dummies for common currency, Free Trade Agreement, common border, common language, common

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<sup>8</sup> Price variables are excluded for statistical estimation simplicity but will be indirectly considered when discussing below the introduction multilateral resistance terms.

colonizer, and past colonial dummy) are included as they are also somehow related to the distance ( $d_{ni}$ ) concept in EK02.

Regression specifications in most related studies include T-variables such as institutional development, educational attainment, trade policy openness, and infrastructure development.<sup>9</sup> These four variables appear significantly (though not robustly) associated with diversification, sophistication, and complexity in several studies (for example Hausmann and others, 2006; Weldemicael, 2012; Ding and Hadzi-Vaskov, 2017), including through Bayesian identification (Giri and others, 2019).

The analysis in this paper also identifies these four variables as the most economically and statistically significant in fostering diversification and complexity relative to other T-variables. As highlighted in the gravity equation literature, exports are not only determined by policy variables of the exporting country, but also by policies of the importing country and therefore they are also included in regressions below. A labor market flexibility variable is introduced to partly capture the wage variable in EK02, assuming rigid labor markets inflate wage costs.

GDP per capita is added as an independent variable acknowledging that it can also approximate wage costs, but mainly to control for potential endogeneity between NHM exports per capita and T-variables. Higher NHM exports can foster GDP and higher GDP can help strengthening T-variables (for example, higher output can facilitate/finance higher educational attainment). Note though that GDP per capita is not included in the calculation of goodness of fit when estimating the predictive power of policy variables.

For estimation purposes EKO2, as well as many other international trade models, can be transformed into a gravity equation specification (Arkolakis and others, 2012).<sup>10</sup> Therefore, some components are added to equation (6) to align it with a gravity equation specification, including: (i) GDP of the source country ( $X_{i,t}$ ); (ii) a remoteness index variable of the importer as a reduced-form control for inward multilateral resistance ( $R_{n,t}$ ); (iii) a remoteness index variable of the exporter as a reduced-form control for outward multilateral resistance ( $R_{i,t}$ ); and (iv) T-variables of the importing country ( $T_{n,t}$ ):<sup>11</sup>

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<sup>9</sup> Trade policy openness and transport infrastructure can be alternatively considered proxies for effective distance between countries.

<sup>10</sup> Yotov and others (2012) provide a comprehensive discussion of gravity equation estimation challenges and techniques.

<sup>11</sup> Although including log GDP per capita and log GDP of the source country can add multicollinearity to the specification (the correlation among them in the sample is 0.5) this does not directly impact the main conclusions of this paper because: (i) the estimated coefficient of log GDP, which is used to calculate the equation (9), appears stable and statistically significant across regressions; (ii) conclusions based on equation (9) do not change if the coefficient is replaced by 1 (as commonly assumed in the gravity equation literature); main conclusions of the paper are not directly based on the coefficient of log GDP per capita.

$$(7) \quad X_{ni,t} = \alpha_1 \log(\gamma) + \alpha_2 \log(X_{n,t}) + \alpha_3 \log(d_{ni}) + \alpha_4 \log T_{i,t} + \alpha_5 \log(w_{i,t}) + \\ + \alpha_6 \log\left(\left(\frac{X}{N}\right)_{i,t}\right) + \alpha_7 \log(X_{i,t}) + \alpha_8 \log(RI_{n,t}) + \alpha_9 \log(RI_{i,t}) + \alpha_{10} \log(T_{n,t}) + \vartheta_{ni} + \\ \mu_t + \varepsilon_{ni,t}$$

As in Wei (1996) and Anderson and van Wincoop (2003), the remote index of a country  $i$  is calculated as a weighted average of the distance of country  $i$  to its trading partners ( $n$ ), where the weights are the incomes of trading partners ( $X_n$ ):<sup>12</sup>

$$(8) \quad RI_i = \sum X_n d_{ni}$$

Since country-pair fixed effects terms are fully correlated with time invariant variables, equation (7) is estimated using the Hausman and Taylor (1981) instrumental variable technique.<sup>13</sup>

After estimating the elasticities of gravity and policy variables within the bilateral trade framework of equation (7), it proves quite insightful to estimate their predictive power at the export country level. The model-predicted level of targeted exports of each country can hence be compared to their actual level to provide a sense of goodness of fit. A complication in doing this is aggregating gravity-equation related variables from the bilateral to the export country level in an intuitive manner. A simple route is to build a Proximity to Markets ( $PM$ ) index for each country, using  $X_n$  and  $d_{ni}$  of equation 7 and their coefficients:

$$(9) \quad \widehat{PM}_i = \sum \frac{X_n^{\hat{\alpha}_2}}{d_{ni}^{\hat{\alpha}_3}}$$

As defined here the  $PM$  index of country  $i$  aggregates the economic size of all trading partners weighted by the inverse of their distance to country  $i$ . Intuitively,  $PM$  is a measure of the trade-generating *gravity pull* that each country experiences. It is clearly an inverse concept of  $RI$  defined in equation (8).

Predicted exports per capita of each country ( $X_i/N_i$ ) are first calculated based only on the  $PM$ 's "gravity pull":

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<sup>12</sup> The use of remoteness indices to control for multilateral resistance could be considered a second-best option, as Yotov and others (2012) suggests that a first-best option would be the use of exporter-time and importer-time fixed effects. However, this potentially first-best option would wash out the identification of the hypothesized determinants of the dependent variables in this study due to multicollinearity with the proposed fixed effects.

<sup>13</sup> Because the dependent variable in equation (3) is specified in logarithmic terms, zero trade flows cannot be included and therefore the information they could provide is neglected. While regressions on goods exports have strong statistical fit despite this omission, regressions on the smaller and less comprehensive set of service exports observations may be more affected by this omission. Yotov and others (2012) indeed suggests that the omission of zero trade flows is more significant for sectoral service trade due to their highly localized consumption and highly specialized production.

$$(10) \log\left(\frac{x_{i,t}}{N_{i,t}}\right) = \beta_0 + \log(\widehat{PM}_{i,t})$$

and then based on  $PM$ ,  $T$  and  $w$  variables, using their coefficients estimated in equation (7), which controlled for potential endogeneity through the inclusion of GDP per capita:

$$(11) \log\left(\frac{x_{i,t}}{N_{i,t}}\right) = \beta_1 + \log(\widehat{PM}_{i,t}) + (\hat{\alpha}_4 \log T_{i,t} + \hat{\alpha}_5 \log(w_{i,t}))$$

As indicated earlier the dependent variables of interest in this paper are those types of exports that are conducive to diversification and/or are identified as *superior* in some studies. These are, specifically, exports of NHM, manufacturing, complex, and service products. Based on SITC Revision 2 classification, the NHM group excludes codes 3000-4999, 6772-6999, and 9000-9999, while manufactures includes codes 6900-8999. Services excludes government services (code 11 under classification EBOPS 2002). Complex products include goods that in Hausmann and others (2013) have a PCI above zero (approximately the top half of the PCI ranking).<sup>14</sup> A group including only HM exports is also analyzed for comparison purposes. The data source for exports of goods and services is UN's Comtrade (United Nations, 2020), which includes the Extended Balance of Payments Services (EBOPS) database.

Gravity equation variables are extracted from the CEPII gravity database constructed by Head and others (2010) and Rose (2004).<sup>15</sup> Variables measuring political stability and governance are extracted from Polity IV (2014) and World Bank (2020a), respectively. Educational attainment data was retrieved from the United Nations Education index (UNDP, 2020) and Barro-Lee (Barro and Lee, 2013). Tariff data comes from the World Integrated Trade Solution (World Bank, 2020b). Infrastructure quality and other measures of T-variables come from the Global Competitiveness Report (World Economic Forum and Harvard University, 2020) and the Doing Business database (World Bank, 2020c). Labor market flexibility is approximated through related subindices in the Global Competitiveness Report and in International Monetary Fund (2019).

#### IV. DESCRIPTION OF KEY VARIABLES

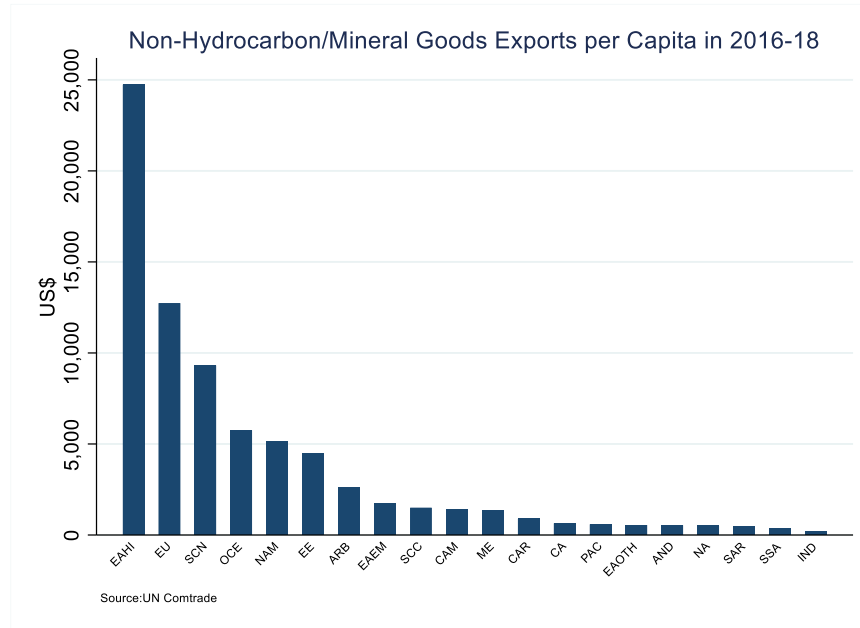
Interesting patterns stand out when looking at per capita levels of the export categories associated with diversification and *superior* exports. In a regional comparison, the level of NHM exports per capita appears associated with GDP per capita (see figure below). Even though some Eastern European (EE) and East Asian Emerging Market (EAEM) countries are commonly considered champions of diversification because of their low concentration indices, the higher income, natural resource abundant countries in Scandinavia (SCN) and Oceania (OCE, Australia and New Zealand) have considerably higher levels of NHM exports

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<sup>14</sup> The Product Complexity Index is available at < <https://atlas.cid.harvard.edu/rankings/product> >.

<sup>15</sup> Available at [http://www.cepii.fr/cepii/en/bdd\\_modele/bdd.asp](http://www.cepii.fr/cepii/en/bdd_modele/bdd.asp).

per capita<sup>16</sup>. And notwithstanding their higher export concentration indices, Southern Cone (SCC) countries (Argentina, Brazil, Chile, Paraguay, and Uruguay) have in average a similar level of NHM exports per capita as EAEM and Central American countries (CAM). This confirms the discussion above suggesting that concentration indices are not an accurate reflection of the development of NHM exports.



But while there is a close relation between income and NHM exports the association is far from perfect. For instance, although natural resource abundant SCN and OCE countries have a higher income per capita than EU countries, their NHM exports per capita are significantly lower. Also prominent is the superlative level of NHM exports per capita of East Asian High-Income economies (EAHI, Hong Kong, Japan, Singapore, South Korea, and Taiwan), even though their income per capita is not considerably different from that of EU and SCN countries.

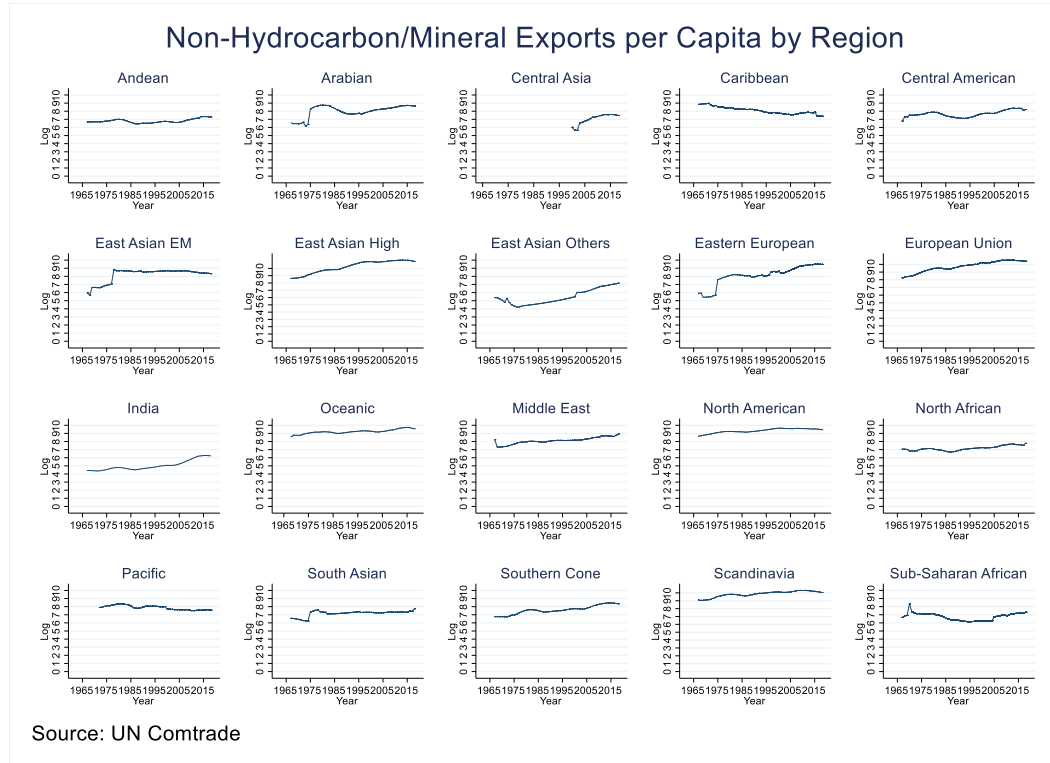
At least two characteristics seem to explain these misalignments between income per capita and NHM exports per capita. One is that, as discussed earlier, even if the production of HM exports is not particularly labor-intensive, the share of labor it demands directly is not negligible, and the share of labor employed in industries and services that support HM production is significant. With less labor force available to other sectors, NHM exports per capita is expected to be lower in countries with significant HM exports. This then could partly explain why HM intensive OCE and SCN countries have lower NHM exports per capita than other countries with similar income per capita.

A second characteristic that can explain the misalignments is that EAHI and other countries that are close to large international markets and therefore participate more intensively in GVCs, have gross NHM exports that overstate their domestic value added. According to the

<sup>16</sup> Regional acronyms are described in Table A.2.

OECD Trade in Value Added (TIVA) database (OECD, 2019), the domestic value added of NHM exports in remote Australia and Chile is 81 and 88 percent of their gross exports, respectively, while the value added in EAHI and EAEM economies is around 60 percent.<sup>17</sup>

The dynamics of NHM exports per capita is also revealing (see panel figure below). In contrast to indices of export diversification and of *superior* exports, NHM exports do not appear affected by commodity prices as they grew resiliently in HM-exporting Andean (AND), SCC, and Sub-Saharan African (SSA) regions throughout the 2000s commodity boom.

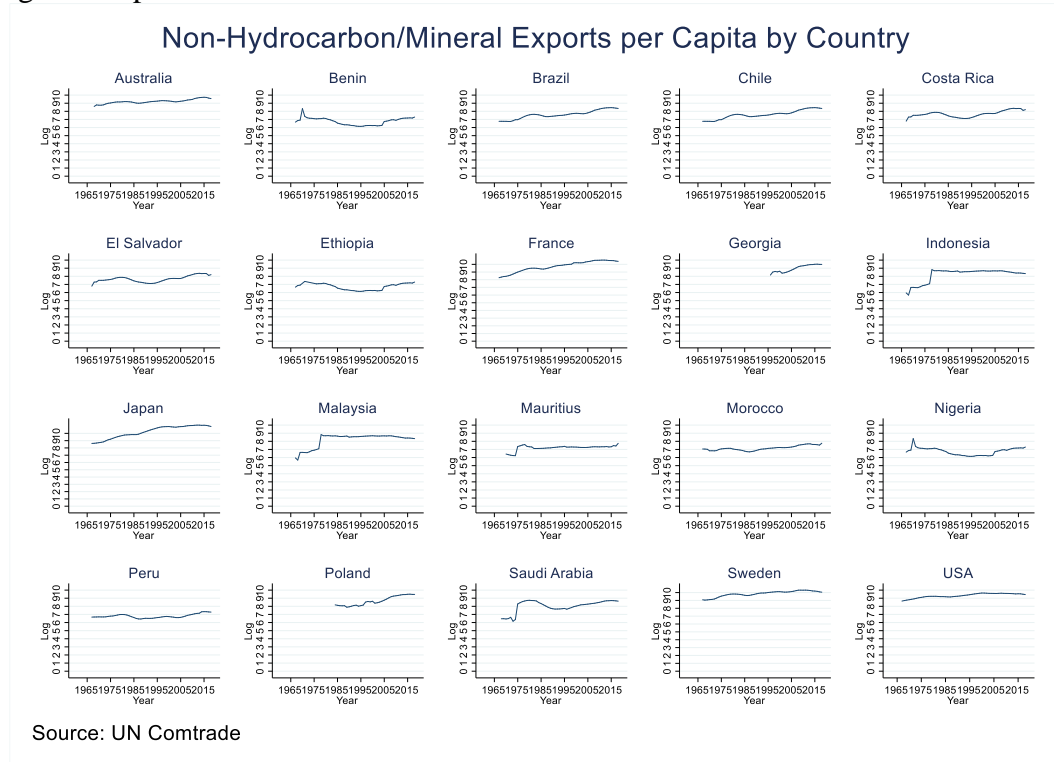


Another important pattern observed in these charts is that NHM exports growth has been faster in regions such as CAM, EE, EAEM, and EAHI, that are closer to the largest world economic centers. East Asian countries exhibited high growth since the 1960s, but their growth decelerated in recent decades. This slowdown does not seem related to a convergence dynamic because EAEM countries decelerated before reaching the level of EAHI countries. NHM exports growth in CAM and EE countries, conversely, accelerated in recent decades, broadly following their economic/trade liberalization episodes. Some remote regions, such as OCE and SCC have also experienced faster NHM exports growth following liberalization episodes.

<sup>17</sup> Data on exports value added is not available for most countries, therefore the rest of the analysis centers on gross exports. Note that all indices of diversification and export “superiority” are subject to this caveat.

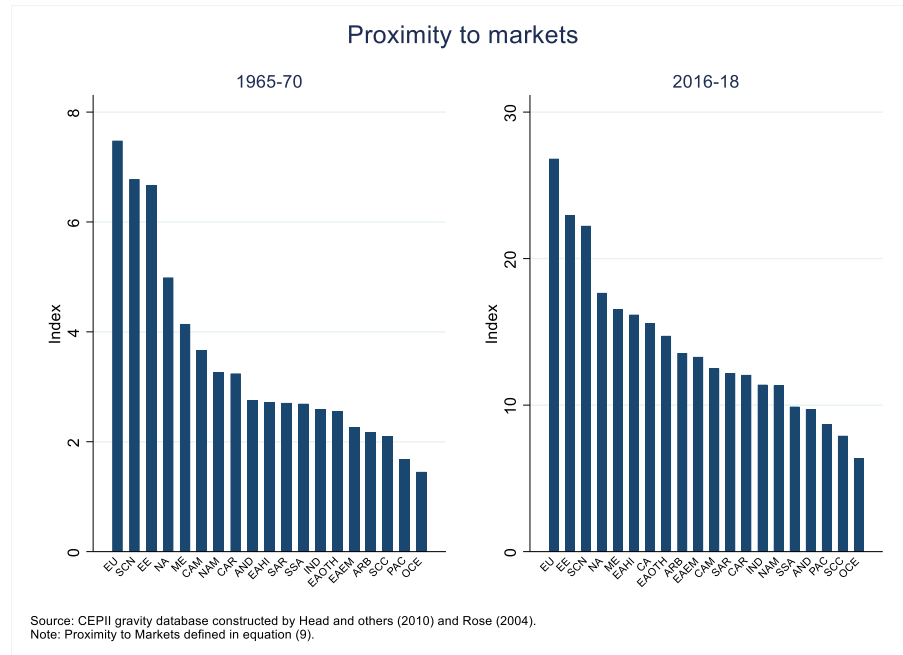


A country-level view of NHM exports growth broadly confirms these regional patterns and provides further insights (see panel figure below). Countries close to large international markets such as Costa Rica, El Salvador, Poland, Malaysia, and Indonesia show the fastest NHM growth, and countries that undertook strong trade liberalization like Costa Rica, Chile, El Salvador, Poland, experienced an acceleration in the aftermath of these reforms. Also remarkable is the stagnation in NHM exports in countries previously considered models of diversification, such as Dominican Republic and Mauritius. These countries though have dynamic services sectors (mainly tourism) that are not included in these charts showing NHM goods exports.



Cross-regional comparisons of the *PM* index confirms that distance to global economic centers appears closely linked to NHM export development. The charts below confirm that regions with fast growth in NHM exports (CAM, EE, EAH, and EAEM) have substantially higher *PM* than most other emerging and developing regions.<sup>18</sup> Remarkably, CAM and EE had very high *PM* already in the 1960s, but as noted earlier, they did not experience a sustained acceleration in NHM exports until they liberalized their trade regimes.

<sup>18</sup> Exponents  $\hat{\alpha}_2$  and  $\hat{\alpha}_3$  in equation (9) are equal to one in these charts.



Already in the 1960s, East Asian countries benefitted from a significantly higher *PM* than many emerging economies. In fact, *PM* of South Korea in 1965 was about 70 percent higher than SCC, a factor that could significantly help explain the often-noted higher export performance of South Korea. The high *PMs* of South Korea and other East Asian countries in the 1960s reflected their proximity to the large and already industrialized Japanese economy, which was not only beneficial as a source of nearby demand but also of capital and technology transfer.

The largeness of the East Asian economic agglomeration and its efficient sea-based interconnection are clearly an advantage relative to the relative isolation of countries in Oceania, Sub-Saharan Africa, and South America.<sup>19</sup> Furthermore, the higher *PM* of East Asian countries has substantially increased in recent decades, as their high *PMs* have fostered their economies and the growth of their economies have further increased their *PMs* (a virtuous circle).

Panel Figure A.1 shows the evolution of T- variables since the 1960s. Many decades ago, EAHI countries already had superior education and infrastructure as well as lower trade barriers than other developing countries, and they improved their relative standing in decades after. This could also help explain their sustained high export performance. Eastern European countries had superlative education in the 1960s, but they improved their relatively weak infrastructure and open their relatively closed trade policies only since the 1980s.

<sup>19</sup> The higher *PM* in East Asia than South America and Sub-Saharan countries is likely higher if estimated based on travel time among countries, not distance among them. East Asian countries are easily and closely connected through the sea, whereas South American and Sub-Saharan African countries are mainly separated by difficult land geographies.

## V. MAIN RESULTS: WHAT EXPLAINS DIVERSIFICATION AND *SUPERIOR* EXPORTS?

### A. Identification of Determinants

Regressions based on equation (7) confirm the importance of proximity/distance to large global markets for export diversification and the development of *superior* exports. The coefficient of distance to partner country in a first regression including only gravity-equation variables (see table below, column 1), indicates that reducing distance by half is associated with a 150 percent increase in NHM exports. Note that the *PM* of the remote SCC and OCE regions is about half of that of CAM, EE, and East Asian regions and therefore the exogenous distance factor on its own can explain a substantially lower level of NHM exports per capita in these remote regions. The point estimate of the coefficient of distance does not change much when including other policy determinants in columns 2-4. The coefficients of most other gravity equation variables are statistically significant and have the expected signs.

Determinants of non-hydrocarbon/mineral (NHM) exports				
Dependent Variable: Log NHM exports	(1)	(2)	(3)	(4)
Log GDP reporter	1.527***	0.660***	0.16	0.825***
Log GDP partner	0.386***	0.22	0.03	0.159*
Log distance	-1.471***	-1.247***	-1.240***	-1.329***
Common currency dummy	0.386***	0.22	0.03	0.159*
Free trade agreement dummy	0.232***	0.177***	0.131*	0.188***
Common border dummy	0.328**	1.971***	1.837***	1.383***
Common language dummy	0.753***	0.619***	0.316**	0.573***
Common colonizer dummy	0.836***	0.253*	0.716***	0.473***
Past colonial link dummy	1.534***	1.338***	1.501***	1.481***
Log GDP per capita	-0.341***	-0.08	0.644***	0.272***
Governance (WB Index)		0.457***	0.176**	
Education (UN Index)		4.854***	3.550***	
Infrastructure (GCR Index)		0.169***	0.283***	
Average Tariff		-0.0299***	-0.0429***	
Labor market flexibility (GCR Index)		0.021		
Labor market flexibility (IMF Index)			2.861***	
Political stability (Polity IV)				-0.0177***
Education (Barro Lee)				0.0438***
Infrastructure (WDI)				0.0238***
Trade liberalization (Wacziarg and Welch, 2003)				0.517***
Constant	-21.94***	0.48	15.44***	-4.178*
Observations	169,968	44,989	31,253	90,352
Rho	0.76	0.94	0.96	0.86

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Panel regressions based on Hausman and Taylor (1981) technique with groups consisting of all combinations of reporter and partner countries in UN Comtrade database. Observations are non-overlapping 5-year averages within the 1962-2018 period, depending on data availability. Regression specification based on equation (7). Multilateral resistance terms and partner country's policy variables included (coefficients not reported). Dependent variable is the logarithm of the value of exports excluding hydrocarbon and mineral products (SITC2 codes 0-2999, 4000-6772, 6900-8999). Trade liberalization dummy based on identification of trade liberalization episodes in Wacziarg and Welch (2008), with a value of 1 if the country has liberalized.

The statistical and economic significance of T- variables in the second column regression is also remarkable, particularly of education. A one standard deviation increase in educational attainment is associated with a 170 percent increase in NHM exports, while same increases in governance and infrastructure quality increase them by also significant 65 and 20 percent, respectively. Consistent with the noticeable acceleration in NHM exports after trade liberalization episodes shown above, reducing the average import tariff from 15 to 5 percent is associated with a significant 45 percent increase in NHM exports. Labor market flexibility is not significant when measured by the Global Competitiveness Report subindex, but a one

standard deviation increase in the IMF labor subindex appears associated with a 45 percent increase in NHM exports in the third column regression.<sup>20</sup>

The fourth regression substantially extends the sample back in time by using the Polity IV political stability index as a proxy for governance; an index of railroad, phone lines, and electricity coverage as a proxy for infrastructure; and, as a proxy for trade policy, a binary variable with value of one if the trade policy regime has been liberalized according to Wacziarg and Welch (2008). In addition, the Barro-Lee education attainment indicator replaces the UN indicator and is interpolated in missing years. The estimated coefficients confirm the importance of these variables except for political stability. The coefficient of the liberalization dummy implies that liberalization episodes in Wacziarg and Welch (2008) are associated with an 80 percent increase in NHM exports per capita.

The significance of these policy determinants remains robust to the inclusion of other T-variables in the Doing Business and Global Competitiveness reports (see table below). Doing Business variables are statistically significant, but estimated coefficients imply only single-digit percent impact of a standard deviation change on NHM exports. As expected, a one standard deviation increase in the overall Doing Business score appears more effective, as it is associated with a 15 percent increase of NHM exports. Among variables in the Global Competitiveness Index, technological readiness has a high magnitude and statistical significance, implying that a standard deviation increase in this variable is associated with a 10 percent increase in NHM exports.

Other determinants of non-hydrocarbon/mineral (NHM) exports				
Dependent Variable: Log non-hydrocarbon/mineral exports	(1)	(2)	(3)	(4)
Log distance	-1.247***	-1.265***	-1.237***	-1.223***
Governance (WB Index)	0.457***	0.391***	0.339***	0.353***
Education (UN Index)	4.854***	4.526***	4.402***	4.528***
Infrastructure (GCR Index)	0.169***	0.157***	0.0809**	0.136***
Average tariff	-0.0299***	-0.0255***	-0.0241***	-0.0246***
Labor market flexibility (GCR Index)	0.02	0.0363	0.0084	0.0390
Construction permits (Doing Business)		0.00	0.00	
Getting credit (Doing Business)		0.00177**	0.00158*	
Investment protection (Doing Business)		0.00288**	0.00317**	
Paying taxes (Doing Business)		0.00	0.00	
Contract enforcement (Doing Business)		0.00	0.00	
Resolving insolvency (Doing Business)		0.00178*	0.00194**	
Macroeconomic environment (GCR Index)			-0.0332*	
Technological readiness (GCR Index)			0.0976***	
Business sophistication (GCR Index)			0.11	
Innovation (GCR Index)			0.08	
Doing Business Score				0.0155***
Constant	0.48	-2.86	-2.20	-9.129**
Observations	44989	44704	44704	44110
Rho	0.94	0.95	0.95	0.95

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Panel regressions based on Hausman and Taylor (1981) technique with groups consisting of all combinations of reporter and partner countries in UN COMTRADE database. Observations are non-overlapping 5-year averages within the 1965-2017 period, depending on data availability. Regression specification based on equation (3). Gravity equation variables, multilateral resistance terms and partner country's policy variables included but coefficients not reported.

<sup>20</sup> Table A.3 shows that coefficients of policy variables the regression in column 2 remain economically and statistically significant to several regression specifications.

Within the broad policy areas of governance, education, and infrastructure there are some specific areas more strongly associated with NHM exports development (see table below). Very prominently, one standard deviation increases in government effectiveness and control of corruption are respectively associated with 60 and 40 percent higher NHM exports. Among education levels, secondary and tertiary education are statistically more significant than primary, while ports and electricity are the most statistically significant infrastructure subcomponents.

Sub-determinants of non-hydrocarbon/mineral (NHM) exports			
Dependent Variable: Log non-hydrocarbon/mineral exports	(1)	(2)	(3)
Log distance	-1.272***	-1.412***	-1.228***
Governance (WB Index)		0.514***	0.277***
Education (UN Index)	4.663***		4.127***
Infrastructure (GCR Index)	0.104***	0.182***	
Average tariff	-0.0267***	-0.0360***	-0.0345***
Labor market flexibility (GCR Index)	0.01	0.05	-0.03
Government effectiveness (WB Index)	0.376***		
Control of corruption (WB Index)	0.520***		
Primary education (GCR Index)		0.002	
Secondary education (GCR Index)		0.00495***	
Tertiary education (GCR Index)		0.00615***	
Port infrastructure (GCR Index)			0.0558**
Electricity infrastructure (GCR Index)			0.0717***
Constant	-1.40 (-0.44)	0.93 -0.30	-10.20** (-3.10)
Observations	44989	44989	44764
Rho	0.94	0.94	0.95

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Panel regressions based on Hausman and Taylor (1981) technique with groups consisting of all combinations of reporter and partner countries in UN Comtrade database. Observations are non-overlapping 5-year averages within the 1962-2018 period, depending on data availability. Regression specification based on equation (7). Multilateral resistance terms and partner country's policy variables included (coefficients not reported). Dependent variable is the logarithm of the value of exports excluding hydrocarbon and mineral products (SITC2 codes 0-2999, 4000-6772, 6900-8999).

Independent variables in equation (7) are also strongly associated with other export groups linked to diversification and complexity (see table below). Distance to partner country has a similar large effect on NHM, complex, and manufacturing exports. Expectedly and very importantly, distance appears less relevant for service exports, thus implying that remote countries could more easily promote diversification by fostering service rather than goods exports. Governance and education are important in fostering all four groups of exports conducive to diversification, but as also expected, overall infrastructure and tariffs are not important for services exports development.<sup>21</sup> Labor market flexibility is, not surprisingly, more significant in determining complex and manufacturing exports than NHM exports.<sup>22</sup> Except for infrastructure development, T-variables are not significantly associated with

<sup>21</sup> Most sub-components of the infrastructure pillar of the Global Competitiveness Report are related to transportation of goods and therefore not as critical to export of services.

<sup>22</sup> NHM exports include some natural resource-based products (for example, agriculture, forestry, and fisheries) and therefore unit labor costs are likely less centrally determinant for them than for manufacturing.

higher HM exports, likely because these exports are more dependent on resource endowment than on the policy environment.

Other determinants of non-hydrocarbon/mineral (NHM) exports				
Dependent Variable: Log non-hydrocarbon/mineral exports	(1)	(2)	(3)	(4)
Log distance	-1.247***	-1.265***	-1.237***	-1.223***
Governance (WB Index)	0.457***	0.391***	0.339***	0.353***
Education (UN Index)	4.854***	4.526***	4.402***	4.528***
Infrastructure (GCR Index)	0.169***	0.157***	0.0809**	0.136***
Average tariff	-0.0299***	-0.0255***	-0.0241***	-0.0246***
Labor market flexibility (GCR Index)	0.02	0.0363	0.0084	0.0390
Construction permits (Doing Business)		0.00	0.00	
Getting credit (Doing Business)		0.00177**	0.00158*	
Investment protection (Doing Business)		0.00288**	0.00317**	
Paying taxes (Doing Business)		0.00	0.00	
Contract enforcement (Doing Business)		0.00	0.00	
Resolving insolvency (Doing Business)		0.00178*	0.00194**	
Macroeconomic environment (GCR Index)			-0.0332*	
Technological readiness (GCR Index)			0.0976***	
Business sophistication (GCR Index)			0.11	
Innovation (GCR Index)			0.08	
Doing Business Score				0.0155***
Constant	0.48	-2.86	-2.20	-9.129**
Observations	44989	44704	44704	44110
Rho	0.94	0.95	0.95	0.95

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Panel regressions based on Hausman and Taylor (1981) technique with groups consisting of all combinations of reporter and partner countries in UN Comtrade database. Observations are non-overlapping 5-year averages within the 1962-2017 period, depending on data availability. Regression specification based on equation (7). Dependent variable is the logarithm of the value of exports excluding hydrocarbon and mineral products (SITC2 codes 0-2999, 4000-6772, 6900-8999). Gravity equation variables, multilateral resistance terms and partner country's policy variables included (coefficients not reported).

As hinted by stylized facts, regression analysis also shows that the *gravity pull* from being close to larger international markets and an open trade regime appear positively related to NHM exports growth not just to their level. In a simple dynamic regression with growth of NHM exports per capita as dependent variable and *PM*, GDP per capita, and the Wacziarg and Welch (2008) trade liberalization dummy as independent variables, the coefficients of *PM* and trade liberalization are positive and statistically significant. The point estimate of the *PM* coefficient implies that doubling *PM* is associated with a 15 percentage points higher annual growth in NHM exports, while the point estimate of the trade liberalization coefficient implies that these episodes are followed by a 6 percentage points acceleration in NHM exports growth.<sup>23</sup> A similar regression on HM exports per capita growth finds this variable strongly related to trade liberalization, but not as much with *PM*.

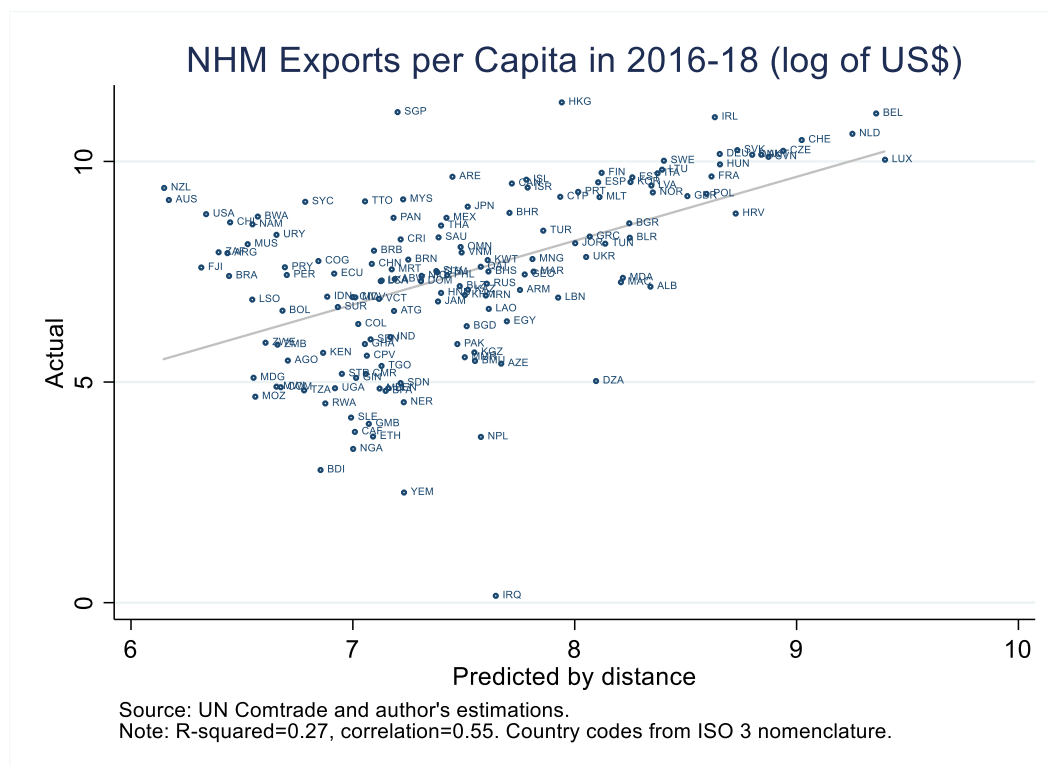
Annual growth of exports on potential determinants.		
Dependent Variable:	NHM exports per capita	HM exports per capita
Lagged growth of dependent variable	0.649	0.012
	0.00	0.233
log(Proximity to Markets)	0.143	0.353
	0.000	0.328
Log(GDP per capita)	-0.013	-0.166
	0.001	0.210
Trade Liberalization dummy	0.059	0.386
	0.000	0.090
Observations	5575	5372

Notes: Results from System GMM regressions. P-values below coefficients. Trade liberalization dummy based on Wacziarg and Welch (2008). Data from UN Comtrade.

<sup>23</sup> The *PM* coefficient implies that the difference in *PM* between Malaysia and Chile (Malaysia's *PM* is 60 percent higher than Chile's) more than explains Malaysia's 2.8 percentage points higher average growth in its NHM exports per capita since the 1960s.

## B. Analysis of Fit: Predictive Power of Proximity and Horizontal Policies

Being located close to large economic centers is a critical advantage for the development of exports conducive to export diversification. This is most evident in a scatter plot (see figure below) comparing actual levels of NHM exports per capita versus those predicted only by *PM*. The correlation between both variables is 0.54 and the R-squared of a fitted line indicates that a quarter of the variation in NHM exports is predicted by countries' exogenous proximity to other economies.<sup>24</sup>



As previous regressions results suggest, the predictive power of *PM* is similarly important for other subcategories of goods exports, except for HM exports, as seen in Panel Figure A.2. Countries that significantly deviate upward from the fitted line in the NHM exports scatter plot similarly deviate in scatter plots of complex and manufacturing exports.<sup>25</sup> Both correlation and R-squared statistics are lower for HM and service exports. Large deviations in the scatter plot of HM exports are related to natural resource abundance, as expected.<sup>26</sup>

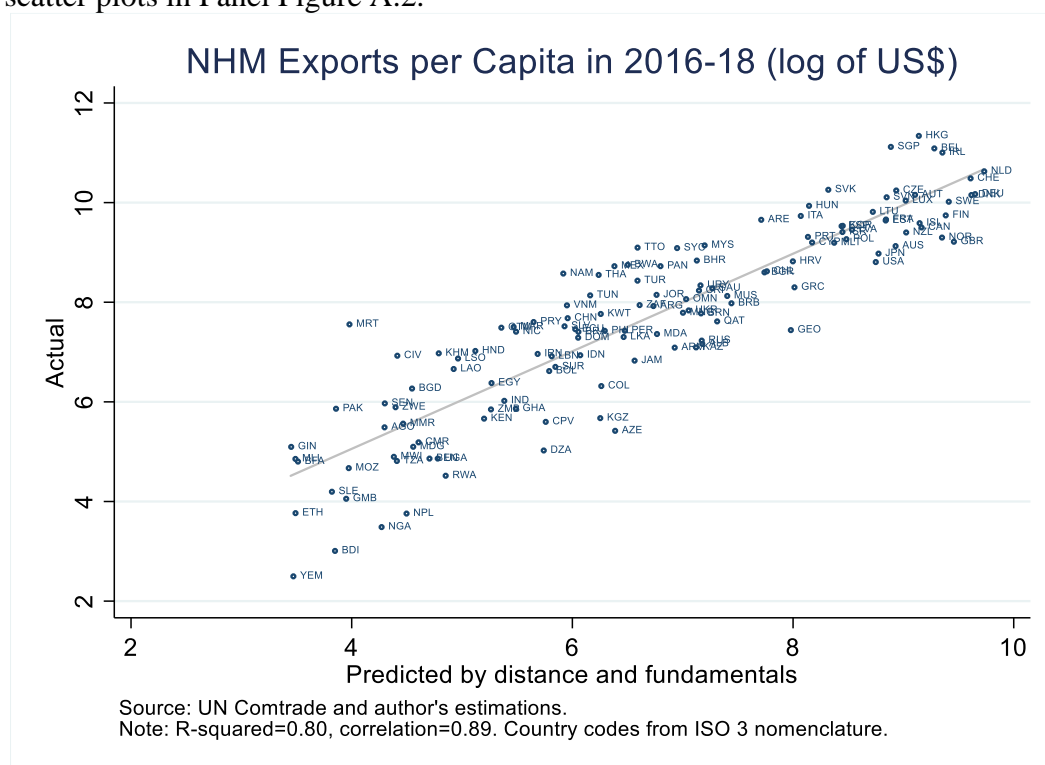
<sup>24</sup> The fitted line has a slope equal to one (a 45-degree line).

<sup>25</sup> The fitted lines of these charts also have slopes equal to one (45-degree lines).

<sup>26</sup> The last two charts in Panel Figure A.2 show that the model's goodness of fit does not change significantly when normalizing NHM exports by workforce instead of population.

Countries that are above their fitted lines very likely have strong diversification-conducive policies that allow them to surpass geographically determined expectations and can therefore hint at “role models” of diversification policies. These include many remote countries, such as Australia (AUS), Chile (CHL), and New Zealand (NZL). The upward deviation of AUS and NZL is remarkably high, higher than that of many well-known role models of diversification, such as Korea (KOR), Japan (JAP), or Malaysia (MYS). Panel Figure A.3 shows deviations from *PM*-predicted NHM and complex exports per capita on a world map.

The predictive power of combining *PM* and policy explanatory variables is impressive, with R-squared and correlation coefficient increasing up to 0.82 and 0.90, respectively.<sup>27</sup> In fact, no country with the *PM* and policy variables of most Sub-Saharan African or Latin American countries has the level of NHM exports per capita of Japan, Korea, or Malaysia. It seems therefore that countries cannot simply bypass the need to strengthen their horizontal institutional, educational, infrastructure, trade, and labor policy framework. This is similarly the case for the development of complex, manufacturing, and service exports as inferred from scatter plots in Panel Figure A.2.<sup>28</sup>



The remarkable improvements in fit for some countries (for example, Chile, OCE, EAH) when adding predictive policy variables suggest that their success is largely associated with their strong horizontal policies. Their policy strength is clearly appreciated in cross regional comparisons in Panel Figures A.4 and A.5.

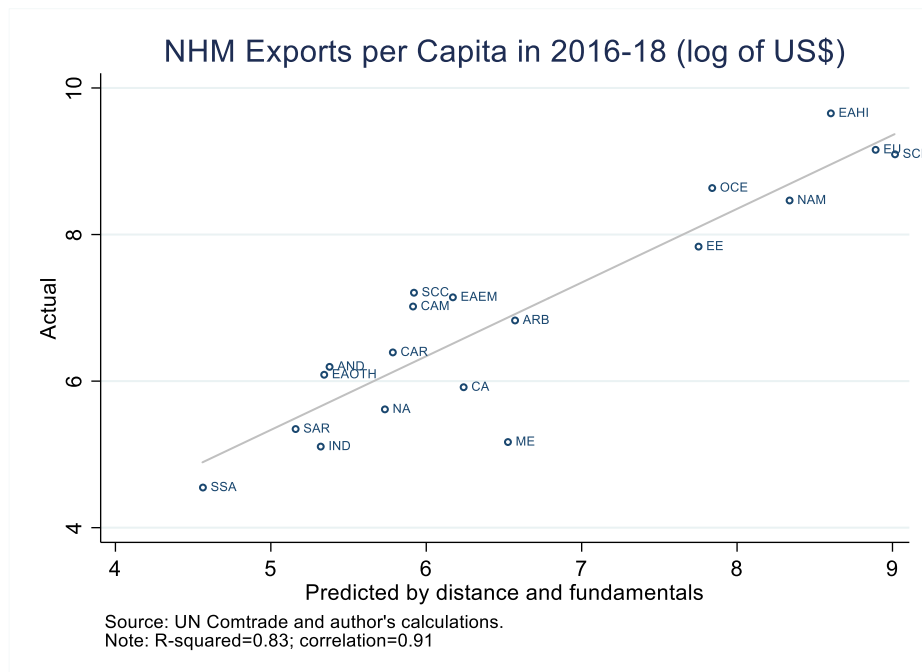
<sup>27</sup> All scatter plots use policy explanatory variables and coefficients from the second regression in page 19.

<sup>28</sup> The slopes of the fitted lines of the charts on NHM, manufacturing, and complex exports are equal to one (45-degree lines).



Remaining outliers after incorporating horizontal policies to the prediction may be the result of vertical diversification policies but some can also reflect less-interesting idiosyncrasies. Positive outliers include countries that participate in GVCs (CAM, East Asia, EE) with high imported components, as noted earlier. Some small outliers are major trading points—including Hong Kong, Panama, and Singapore—with exports that include reexports with low domestic value added. And many outliers are oil exporting countries with highly subsidized industries, which may be economically inefficient and therefore not clearly desirable models of diversification.

Aggregating actual and predicted NHM exports per capita at the regional level further clarifies the big picture. A scatter plot by regions broadly maintains the same R-squared and correlation coefficients as the plot by country. High performing EE countries have NHM exports per capita in line with prediction by *PM* and policy variables identified in this paper. East Asian countries generally export above prediction, but also Central America (CAM) and Southern Cone (SCC) subregions. Sub-Saharan African countries' weak performance is broadly well predicted by its *PM* and policies.



## VI. CONCLUDING REMARKS

This paper shows that to identify the determinants of export diversification it is better to analyze directly export groups associated with diversification instead of focusing on commonly used diversification indices because the latter are extremely sensitive to exogenous HM exports. For the same reason it is more accurate to identify determinants of *superior* exports by focusing on their level instead of commonly used indices. While the

Economic Complexity Index (ECI) aims to be a proxy of the *productive capabilities* of an economic system, its sensitivity to commodity-related factors, completely unrelated to an economy's productivity, severely distorts its accuracy.

The data generating process of export categories linked to diversification can be modeled with high goodness of fit as a gravity equation with horizontal policy variables. *PM* on its own explains about a quarter of the variation in NHM exports per capita and cutting distance to markets by half increases NHM exports by about 150 percent. The lower sensitivity of service exports suggests that they could provide a more feasible route to diversification for remote countries than goods exports. Interestingly, *PM* is also associated with considerably higher growth rates of NHM and complex exports per capita. Table A.5 describes how identifying determinants of export diversification with the methodology proposed in this paper leads to more statistically robust estimates than other related studies.

The centrality of *PM* underscores the need to effectively shorten distance to other economies by enhancing connectedness at all levels, reducing trade policy barriers, enhancing trade facilitation, strengthening transport infrastructure, investing in top-notch communication technology (particularly on internet connectivity to support the digital economy), and fostering technological diffusion (including through educational exchange programs). Enhancing connectedness is crucial for the most remote economies, thus better allowing their production agents to more significantly tap from backward and forward linkages to large global economic centers.

Adding policy variables related to the standard EK02 model, most of which have been identified less robustly in other studies, provides remarkably accurate predictions of cross-country variation in NHM, complex, and manufacturing exports per capita. Such strong fits indicate that no country has the level of NHM exports per capita of East Asian high-performing economies if they have the *PM* and quality of horizontal policy variables normally seen in Latin America or Sub-Saharan Africa. The importance of the identified policy variables is most clearly appreciated in countries that have high NHM exports per capita despite their remoteness (notably Australia, Chile, and New Zealand), as they seem to have been able to significantly offset the natural drawback of distance via strong policy frameworks.

Strengthening horizontal policy areas such as institutions, education, technological readiness, may seem a daunting and extremely long-term task, but significant payoff can be obtained from short-to-medium term reforms that reduce trade policy barriers, strengthen infrastructure, and remove excessively restrictive regulations, including of the labor market. Note also that within the wide objective of governance strengthening, more focused reforms to improve government effectiveness and control corruption seem particularly productive. And concentrating on strengthening port and electricity infrastructure seems most important among all infrastructure areas.

The identified statistical importance of horizontal policies does not necessarily deny some potential contribution from sector-specific policies. One potential reason is that the effective application of vertical policies could be partly correlated with horizontal policy covariates. It

is quite possible that vertical policies are more successfully implemented in countries with higher government effectiveness, better control of corruption, and overall education.

In fact, the identified horizontal policy determinants of diversification could shed light to some second-best sector-specific interventions to offset economy wide policy/institutional weaknesses, such as is sometimes done through Special Economic Zones (SEZs). One such case happened in Mauritius in the 1970s. With high imports restrictions (average tariffs at around 100 percent) and rigid labor market legislation, this country fostered export diversification and complexity by setting up SEZs without import restrictions and with a flexible labor legislation, including one of the lowest minimum wage-to-GDP ratios in the world. These schemes most likely played an important role in Mauritius' notable export development, together with its relatively high overall institutional strength and educational quality.

Countries could similarly strengthen diversification determinants identified in this paper at a sector level without resorting to SEZs. For example, governments could provide technical education to strengthen a specific sector's productivity, develop sector-specific infrastructure, or impose lower restrictions on imported inputs to favored sectors, although this last option is more prone to rent seeking.

Given a long and vast history of sector-specific and industrial policy failures, any assessment of these policies should consider not only their potential effectiveness in promoting targeted exports but also the risk that they could lead to fiscal erosion, facilitate rent seeking, contribute to a "race to the bottom" in international taxation, or weaken multilateralism. In this sense, tax incentives, subsidized credit, exchange rate manipulation, sector-specific trade protection, are among the "riskier" options and there seems to be no strong evidence identifying the effectiveness of this kind of "hard industrial policies" (Harrison and Rodriguez-Claire, 2012).

<sup>29</sup> On the other hand, diversification strategies centered on strengthening horizontal policies are not only less controversial but, as the statistical analysis shown in this study suggests, constitute the backbone of export diversification.

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<sup>29</sup> Harrison and Rodriguez-Claire (2012) reviews empirical work on industrial policy effectiveness and finds no support for "hard" interventions that distort prices. Several surveys have sought to provide evidence on the importance of industrial policies in promoting export diversification, particularly in East Asian countries (see for example, Rodrik, 2004; Rodrik, 2008; or Cherif and Hasanov, 2019). Yet, as Harrison and Rodriguez-Claire (2012) reflects, while it is clear that East Asian countries pursued industrial policies it is not clear that those policies should be credited for their successful industrialization, especially considering that these countries also excelled on horizontal policies, as shown in this WP. In addition, other studies have negatively assessed the role of industrial policies, notably since Caves and Uekesa (1976), which suggested that Japan's MITI created recession cartels and entry barriers that resulted in allocative inefficiency.

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**Table A.1: Determinants of Economic Complexity Index (ECI) and ECI Plus**

Dependent Variable:	ECI		ECI Plus	
Log Complex Exports per Capita	0.077***	0.079***	0.081***	0.083***
Hydrocarbon and Mineral Exports-to-GDP (Sachs and Warner, 1995)		-0.162***		-0.292***
Constant	-0.174***	-0.184***	-0.175***	-0.175***
Observations	4,370	4,207	4,274	4,207
R-squared	0.47	0.44	0.35	0.45

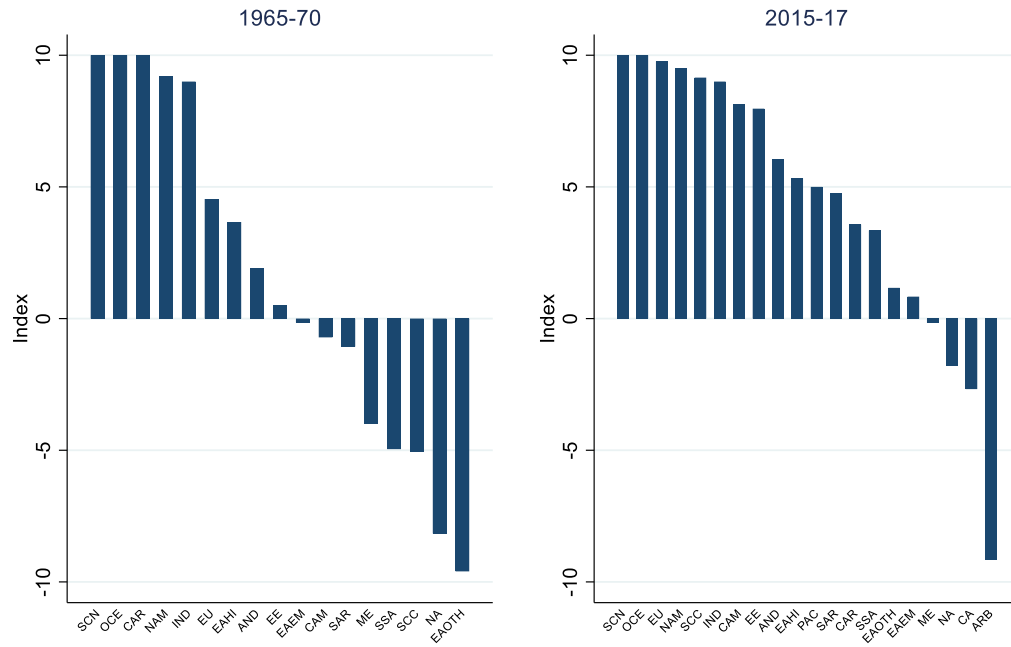
Sources: Hausmann and others (2013), World Economic Outlook (IMF), UN Comtrade, and author's calculations.

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Fixed effects regression.

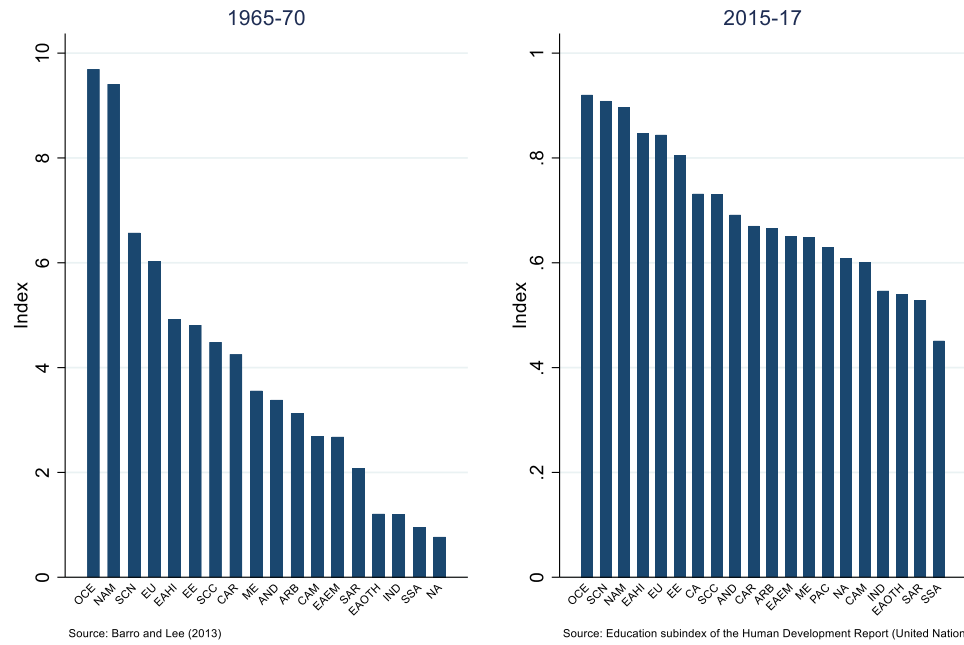


Table A.2: List of Countries by Regional Group

Region	Region Code	Country	Region	Region Code	Country	Region	Region Code	Country
Andean	AND	Bolivia	Eastern Europe	EE	Albania	Pacific Isl.	PAC	Tonga
		Colombia			Bosnia Herz.			Tuvalu
		Ecuador			Bulgaria			Vanuatu
		Peru			Croatia	South Asia	SAR	Afghanistan
		Venezuela			Cyprus			Bangladesh
Arab	ARB	Bahrain			Czechia			Bhutan
		Brunei			Estonia			Nepal
		Kuwait			Georgia			Pakistan
		Oman			Hungary			Sri Lanka
		Qatar			Latvia			Timor-Leste
		Saudi Arabia			Lithuania	Southern Cone	SCC	Argentina
		UAE			Montenegro			Brazil
		Yemen			N. Macedonia			Chile
Central Asia	CA	Armenia			Poland			Paraguay
		Azerbaijan			Moldova			Uruguay
		Belarus			Romania	Scandinavia	SCN	Denmark
		Kazakhstan			Russia			Finland
		Kyrgyzstan			Serbia			Iceland
		Tajikistan			Serb. and Mont.			Norway
		Turkmenistan			Slovakia			Sweden
		Uzbekistan			Slovenia	Sub-Saharan Africa	SSA	Angola
Central Am. & Mexico	CAM	Costa Rica			Turkey			Benin
		El Salvador			Ukraine			Botswana
		Guatemala	European Union	EU	Andorra			Burkina Faso
		Honduras			Austria			Burundi
		Mexico			Belgium			Cabo Verde
		Nicaragua			France			Cameroon
		Panama			Germany			Central African Rep.
Caribbean	CAR	Anguilla			Greece			Chad
		Antig. & Barb.			Greenland			Comoros
		Aruba			Ireland			Congo
		Bahamas			Italy			Cote d'Ivoire
		Barbados			Luxembourg			D.R. Congo
		Belize			Malta			Djibouti
		Bermuda			Netherlands			Eritrea
		Cayman Isds			Portugal			Ethiopia
		Cuba			Spain			Gabon
		Dominica			Switzerland			Gambia
		Dominican Rep.			U.K			Ghana
		French Guiana	India	IND	India			Guinea
		Grenada	Middle East	ME	Iran			Guinea-Bissau
		Guadeloupe			Iraq			Kenya
		Guyana			Israel			Lesotho
		Haiti			Jordan			Liberia
		Jamaica			Lebanon			Madagascar
		Martinique			Syria			Malawi
		Montserrat	North Africa	NA	Algeria			Mali
		St. Kitts & Nevis			Egypt			Mauritania
		Saint Lucia			Libya			Mauritius
		St. Vct. & Gren.			Morocco			Mayotte
		Suriname			Tunisia			Mozambique
		Trinidad & Tob.	North America	NAM	Canada			Namibia
East Asia Emerging	EAEM	China			USA			Niger
		China, Macao SAR	Oceania	OCE	Australia			Nigeria
		Indonesia			New Zealand			Rwanda
		Malaysia	Pacific Isl.	PAC	Cook Isds			Sao Tome & Princ.
		Philippines			FS Micronesia			Senegal
		Thailand			Faeroe Isds			Seychelles
		Viet Nam			Fiji			Sierra Leone
East Asia High Income	EAHI	China, Hong Kong			French Polynesia			Somalia
		Japan			Kiribati			South Africa
		Singapore			Maldives			Sudan
		South Korea			New Caledonia			Togo
East Asia Others	EAOTH	Cambodia			Palau			Uganda
		Lao PDR			Pap. New Gn.			Tanzania
		Mongolia			Samoa			Zambia
		Myanmar			Solomon Isds			Zimbabwe

**Panel Figure A.1: Index of Determinants of Economic Complexity****Political Stability**

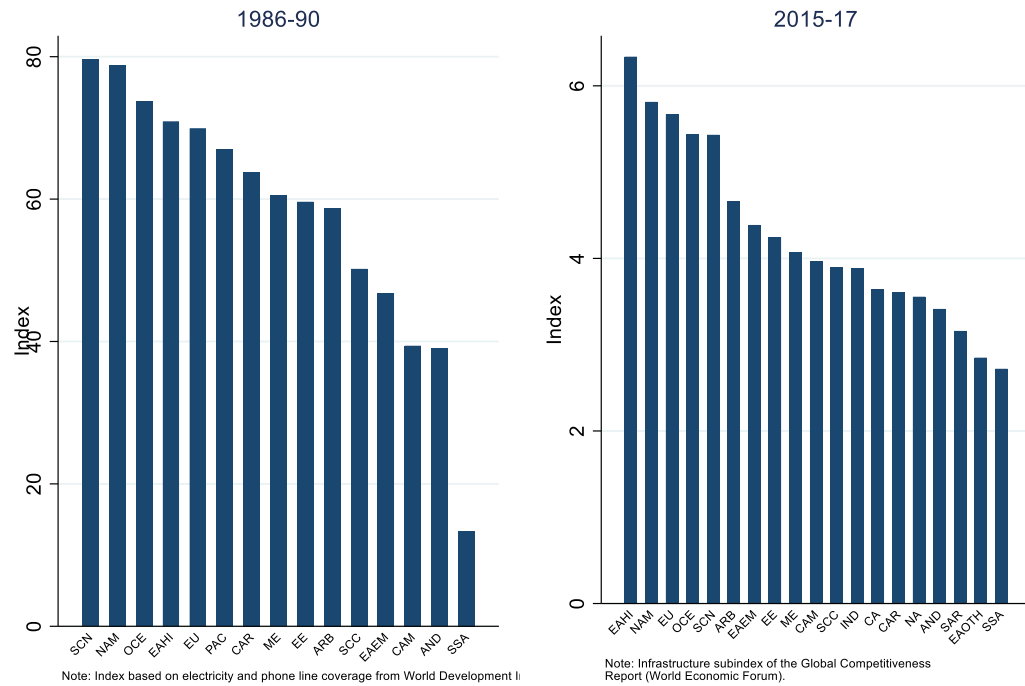
Source: Center for Systemic Peace, Polity IV dataset.

**Educational Attainment**

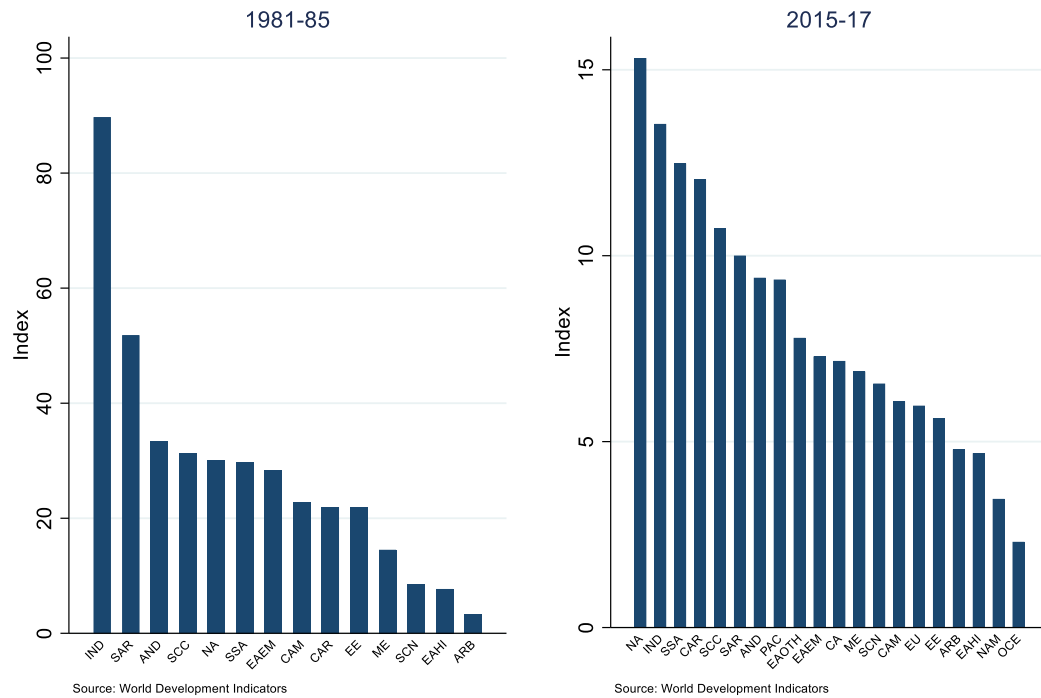
Source: Barro and Lee (2013)

Source: Education subindex of the Human Development Report (United Nations)

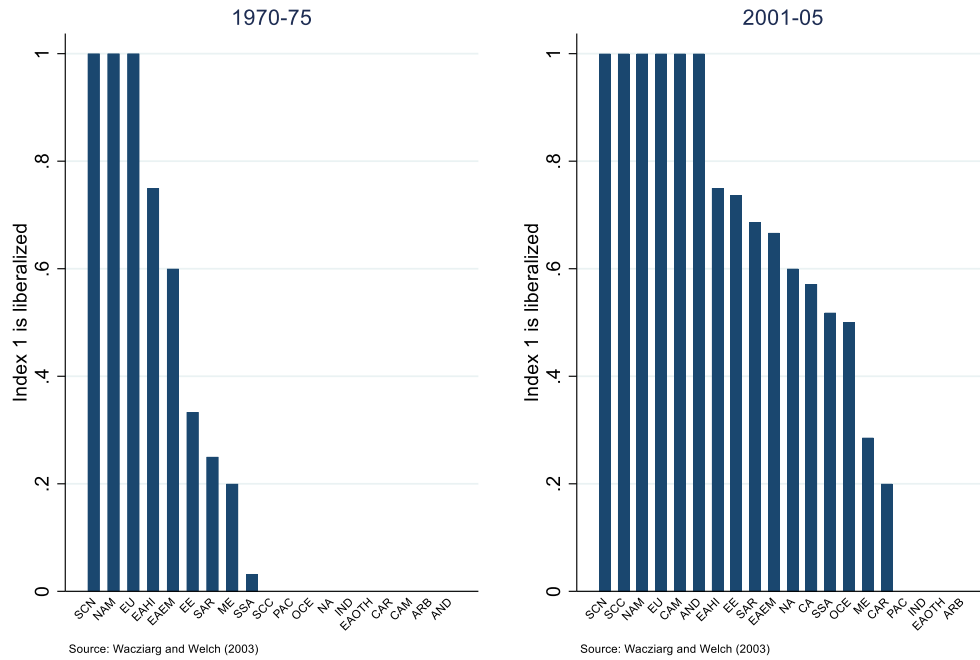
## Infrastructure



## Average Imports Tariffs



# Liberalized Trade



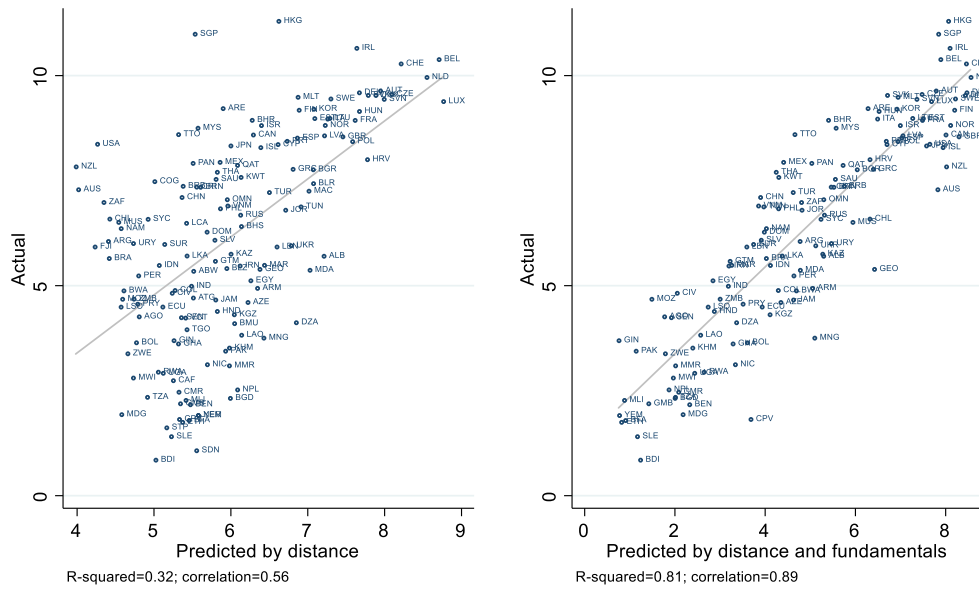
**Table A.3: Determinants of exports by regression specification**

Dependent Variable: Log of non-hydrocarbon/mineral exports							
Regression specification	Hausman-Taylor	Hausman-Taylor	Hausman-Taylor	Pooled OLS	Between Effects	Random Effects	Fixed Effects
Log GDP reporter	0.756***	0.484***	0.756***	1.354***	1.345***	1.369***	-0.548***
Log GDP partner	0.858***	0.960***	0.858***	0.920***	0.904***	0.925***	0.347***
Log distance	-1.279***	-0.616***	-1.279***	-1.420***	-1.444***	-1.465***	
Common currency dummy	0.22	0.368**	0.22	-0.193*	-0.27	-0.01	
Common border dummy	1.888***	2.999***	1.888***	1.308***	1.327***	1.260***	
Common language dummy	0.617***	0.899***	0.617***	0.676***	0.773***	0.638***	
Common colonizer dummy	0.339**	0.327**	0.339**	0.602***	0.561***	0.444***	
Past colonial link dummy	1.228***	1.309***	1.228***	0.526***	0.482***	0.790***	
Log GDP per capita	-0.15	-0.04	-0.15	-0.918***	-0.973***	-0.680***	0.866***
Governance (WB Index)	0.484***	0.422***	0.484***	0.634***	0.505***	0.789***	0.099*
Education (UN Index)	5.099***	3.924***	5.099***	1.166***	1.498***	0.799***	1.031**
Infrastructure (GCR Index)	0.175***	0.166***	0.175***	0.694***	0.864***	0.307***	0.113***
Average Tariff	-0.0310***	-0.0304***	-0.0310***	-0.0722***	-0.0784***	-0.0421***	-0.0197***
Labor market flexibility (GCR Index)	0.02	0.04	0.02	-0.334***	-0.374***	-0.04	0.03
Constant	-0.86	-19.34***	-0.86	-11.67***	-12.17***	-11.28***	-7.223**
	(-0.27)	(-14.50)	(-0.27)	(-13.71)	(-8.86)	(-9.76)	(-3.00)
Partner country policy variables	No	Yes	Yes	Yes	Yes	Yes	Yes
Multilateral resistance proxies	No	No	Yes	Yes	Yes	Yes	Yes
Observations	44,989	44,989	44,989	44,989	44,989	44,989	44,989
Rho	0.93	0.94	0.93			0.81	0.94

Notes: \* p<0.1, \*\* p<0.05, \*\*\* p<0.01. Panel regressions based on Hausman and Taylor (1981) technique with groups consisting of all combinations of reporter and partner countries in UN Comtrade database. Observations are non-overlapping 5-year averages within the 1962-2018 period, depending on data availability. Regression specification based on equation (7). Multilateral resistance terms and partner country's policy variables included (coefficients not reported). Dependent variable is the logarithm of the value of exports excluding hydrocarbon and mineral products (SITC2 codes 0-2999, 4000-6772, 6900-8999).

**Panel Figure A.2: Actual vs Predicted Exports**

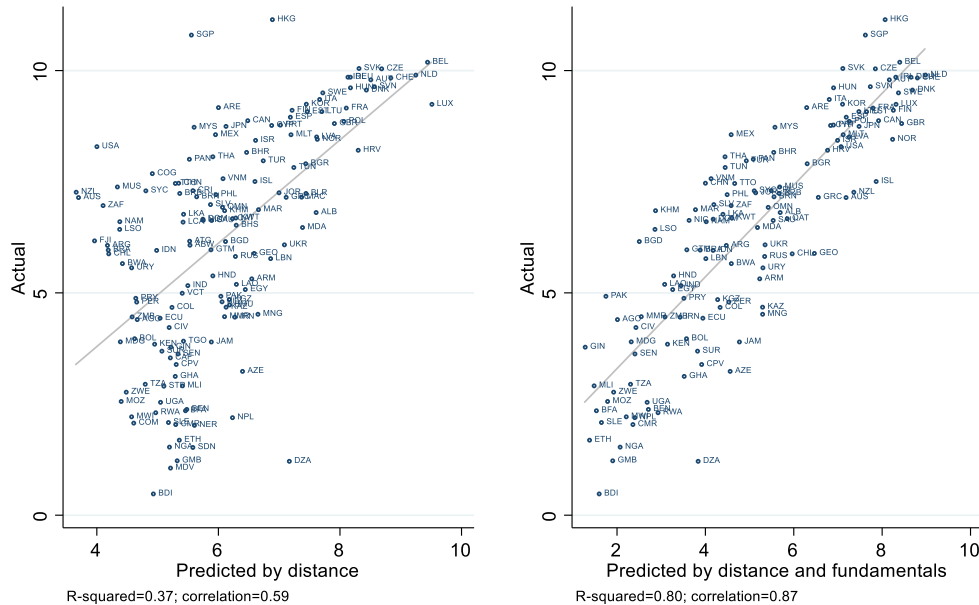
Complex Exports per Capita in 2016-18 (log of US\$)



Source: UN COMTRADE and author's calculations.

Note: Complex exports are those with Product Complexity Index (Hausmann and others, 2013) above zero. Country codes from ISO 3 nomenclature.

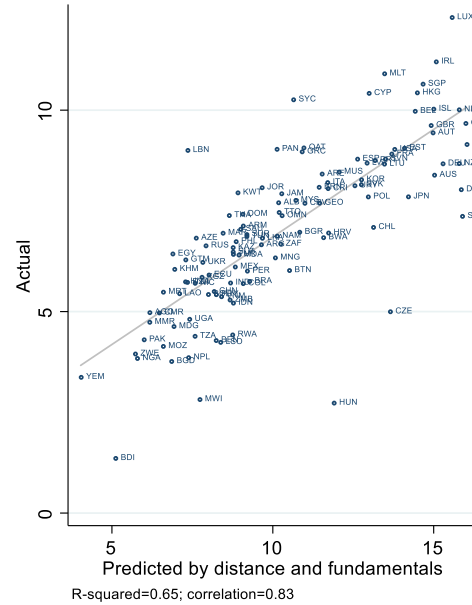
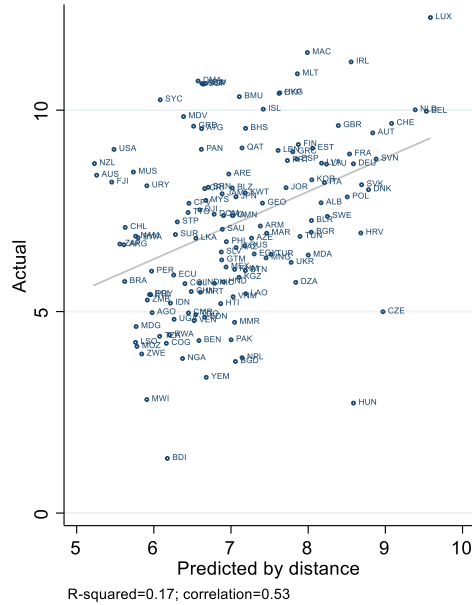
Manufacturing exports per capita in 2016-18 (log of US\$)



Source: UN COMTRADE and author's calculations.

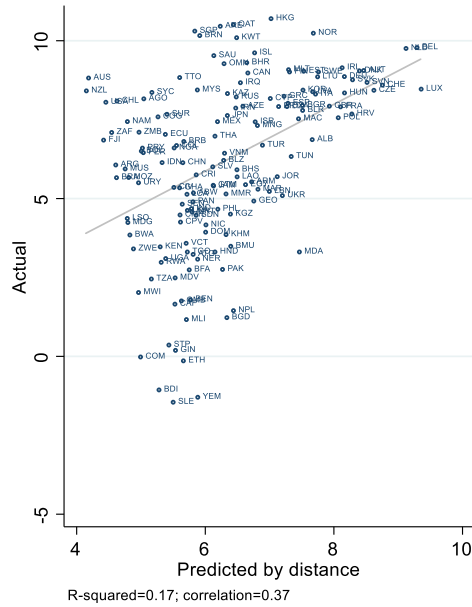
Note: Exports SITC Rev2 from 6900 to 8999. Country codes from ISO 3 nomenclature.

## Service Exports per Capita in 2016-18 (log of US\$)



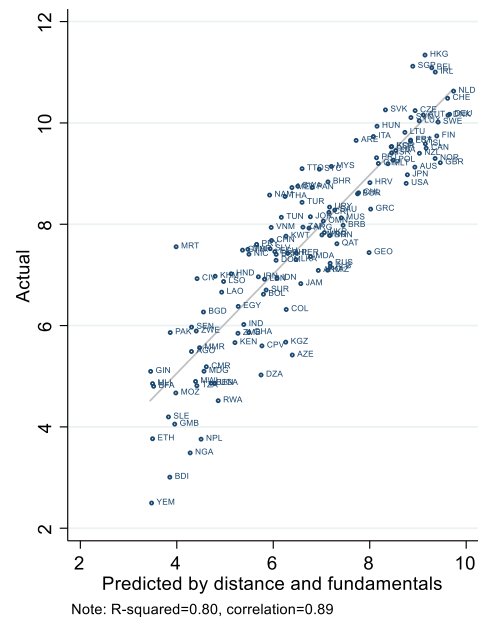
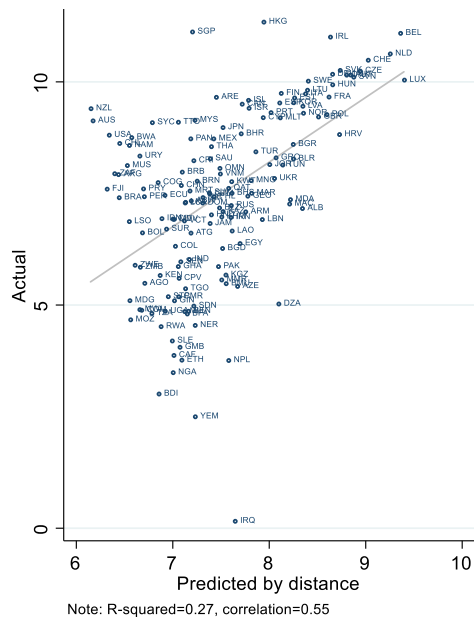
Source: UN COMTRADE and author's calculations.  
Note: Country codes from ISO 3 nomenclature.

## Hydrocarbon and Mineral Exports per Capita in 2016-18 (log of US\$)



Source: UN comtrade and author's calculations.  
Note: Country codes from ISO 3 nomenclature.

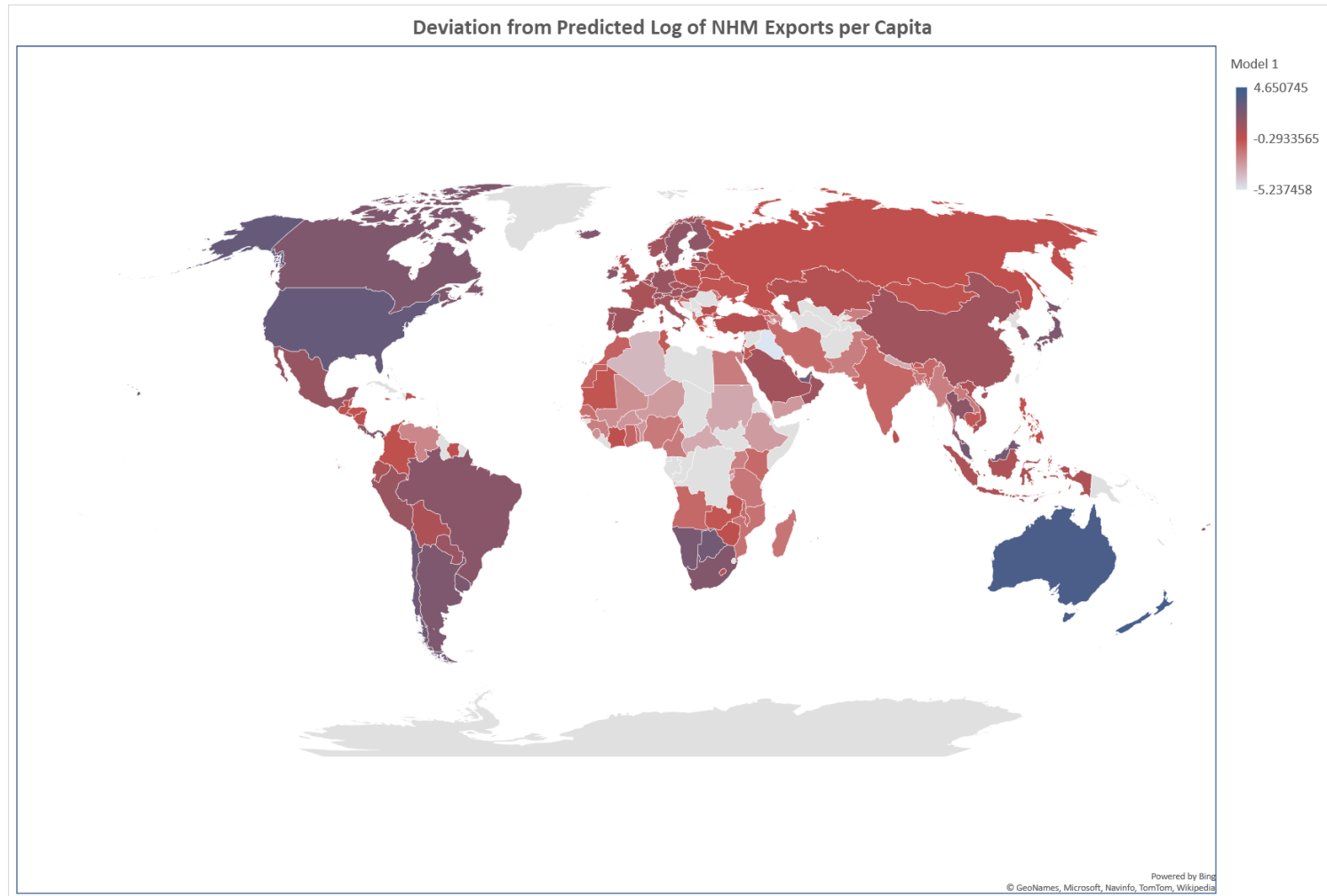
### NHM Exports per Worker in 2016-18 (log of US\$)

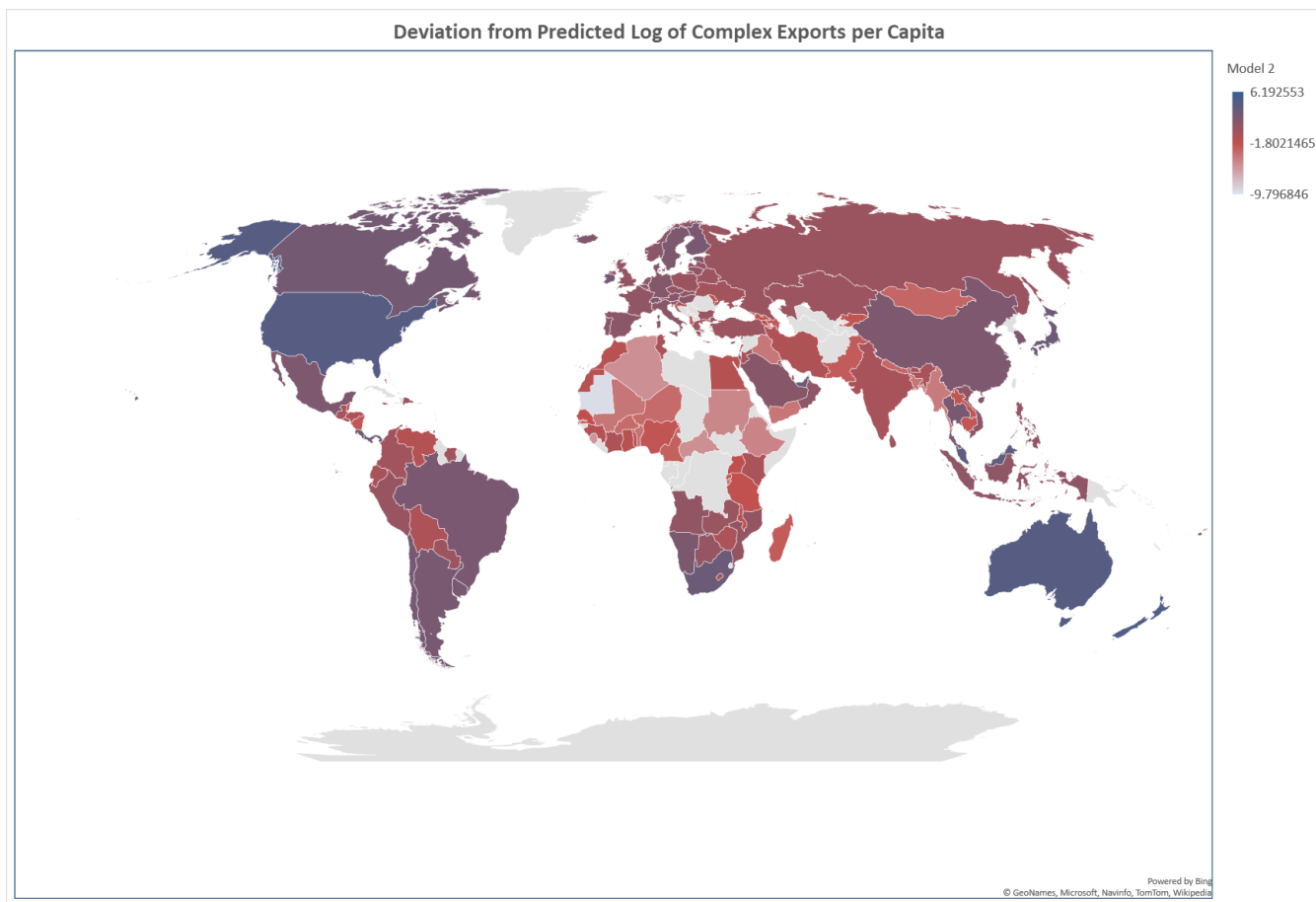


Source: UN Comtrade

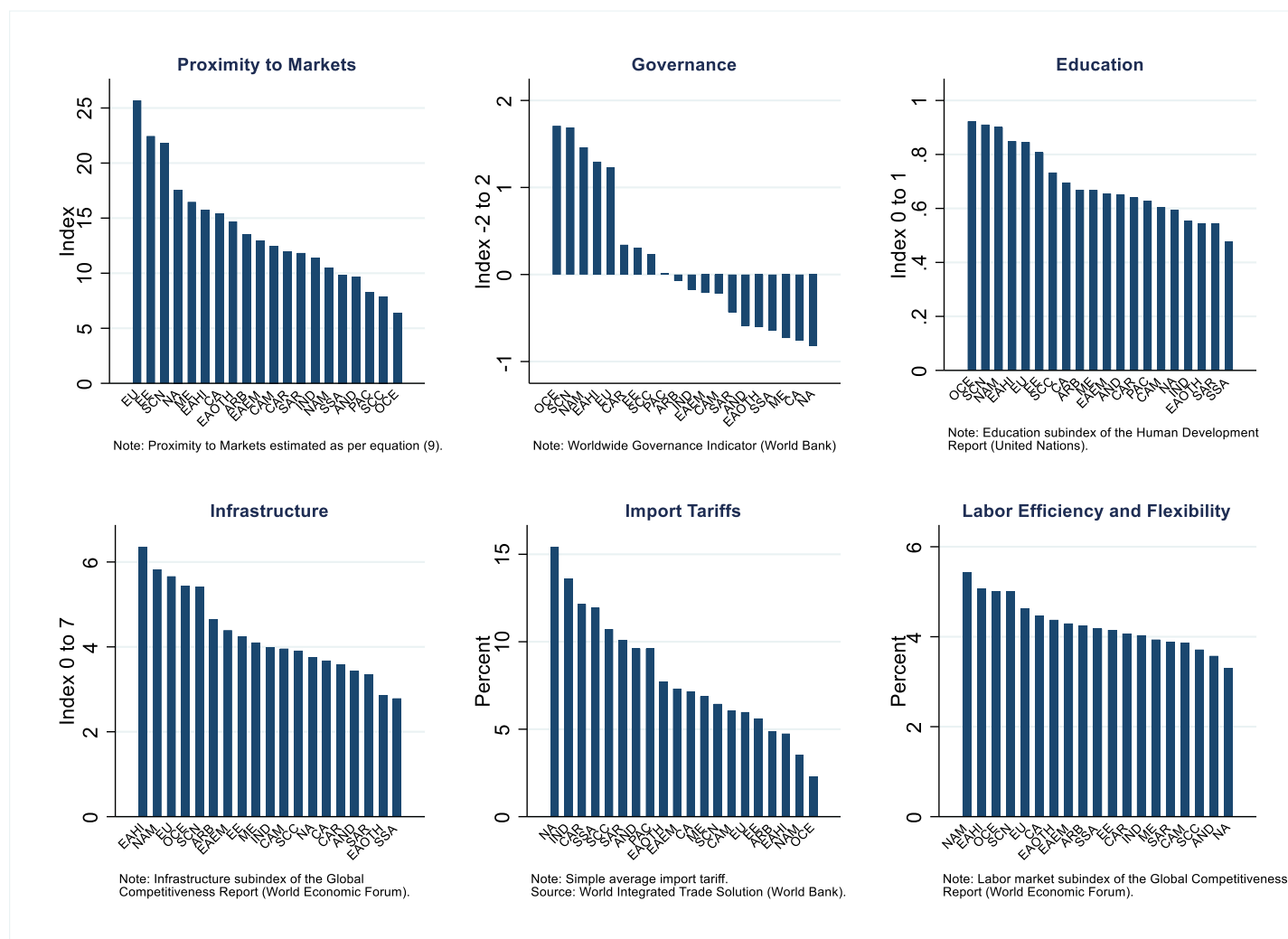


**Panel Figure A.3: Deviation of Actual Exports from Predicted-by Distance Exports**

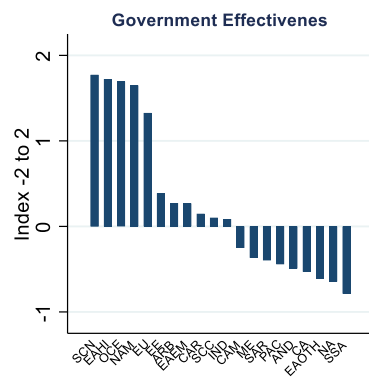




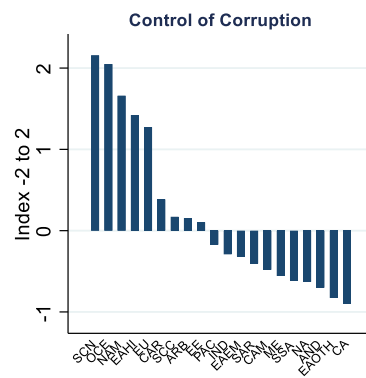
**Panel Figure A.4: Distance and Fundamentals by Subregion in 2015-2017**



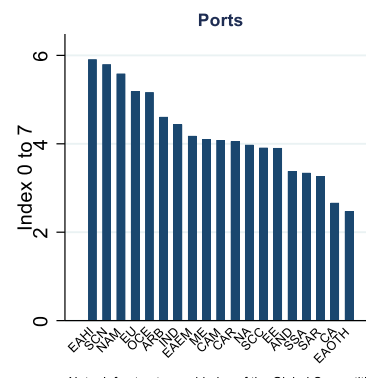
**Panel Figure A.5: Other Explanatory Variables by Subregion in 2015-2017**



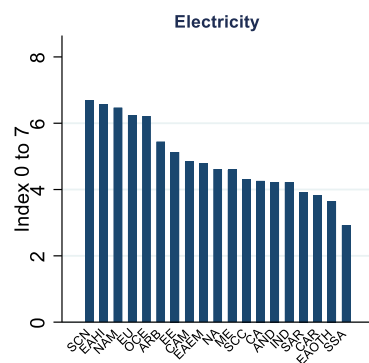
Note: Worldwide Governance Indicator (World Bank)



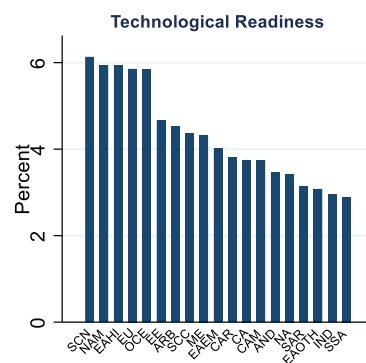
Note: Worldwide Governance Indicator (World Bank)



Note: Infrastructure subindex of the Global Competitiveness Report (World Economic Forum).



Note: Infrastructure subindex of the Global Competitiveness Report (World Economic Forum).



Note: Labor market subindex of the Global Competitiveness Report (World Economic Forum).

**Table A.4: Summary Results by Study**

	This Study Salinas (2021)	Hausmann and others (2009)	Weldemicael (2012)	Ding and Hadzi-Vaskov (2017)	Giri and others (2019)
<i>Dependent Variable:</i>	NHM and Complex Exports per capita	Exports Sophistication Index	Exports Sophistication Index	Exports Concentration, Sophistication, and Complexity Indices	Exports Diversification
<i>Independent Variables:</i>					
Proximity to Markets	Significant in all specifications (robust to inclusion of GDP per capita)	Not included	Significant in all specifications	Not included	Not included
Governance		Not robust to inclusion of GDP per capita	Only significant in sub- sample regressions (all include GDP per capita)	Not included	Significant in some regressions
Education		Not robust to inclusion of GDP per capita		Significant in most specifications	Primary education significant in most regressions
Infrastructure		Not included	Not included	Significant in most specifications	Significant in some regressions
Trade Openness		Not included	Not included	Significant in most specifications	Trade-to-GDP significant in some regressions
Adjusted R-Squared with larger specification	0.80 to 0.86	0.70 to 0.75	0.70 to 0.73	0.457 to 0.751	0.457 to 0.751